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# An exponential cone programming approach for managing electric vehicle charging

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# An Exponential Cone Programming Approach for Managing Electric Vehicle Charging

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To support the rapid growth in global electric vehicle adoption, public charging of electric vehicles is crucial. We study the problem of an electric vehicle charging service provider, which faces (1) stochastic arrival of customers with distinctive arrival and departure times, and energy requirements as well as (2) a total electricity cost including demand charges, costs related to the highest per-period electricity used in a finite horizon. We formulate its problem of scheduling vehicle charging to minimize the expected total cost as a stochastic program (SP). As this SP is large-scale, we solve it using exponential cone program (ECP) approximations. For the SP with unlimited chargers, we derive an ECP as an upper bound and characterize the bound on the gap between their theoretical performances. For the SP with limited chargers, we then extend this ECP by also leveraging the idea from distributionally robust optimization (DRO) of employing an entropic dominance ambiguity set: Instead of using DRO to mitigate distributional ambiguity, we use it to derive an ECP as a tractable upper bound of the SP. We benchmark our ECP approach with sample average approximation (SAA) and a DRO approach using a semi-definite program (SDP) on numerical instances calibrated to real data. As our numerical instances are large-scale, we find that while SDP cannot be solved, ECP scales well and runs efficiently (about 50 times faster than SAA) and consequently results in a lower mean total cost than SAA. We then show that our ECP continues to perform well considering practical implementation issues, including a data-driven setting and an adaptive charging environment. We finally extend our ECP approaches (for both the uncapacitated and capacitated cases) to include the pricing decision and propose an alternating optimization algorithm, which performs better than SAA on our numerical instances. Our method of constructing ECPs can be potentially applicable to approximate more general two-stage linear SPs with fixed recourse. We also use ECP to generate managerial insights for both charging service providers and policymakers.

*Key words* : stochastic programming, exponential cone programming, electric vehicle, demand charge, robust optimization *History* : March 30, 2023.

## 1. Introduction

Electric vehicles (EVs) are considered among the most promising technologies to decarbonize the transportation sector, with market share expected to grow from 1% in 2018 to about 30% of vehicle sales worldwide by 2030 (J.P. Morgan 2018). A key to the mass adoption of EVs is the ease of charging, where public charging will play an increasingly important role, e.g., in supporting adopters without residential charging as well as reducing "range anxiety" (Wood et al. 2017, McKinsey & Company 2018). Examples of public charging service providers include EVgo, Tesla, and ChargePoint, which have the largest market shares in the U.S. It is estimated that this charging market in the U.S. will grow to \$18.6 billion, and the number of charging points will expand to 40 million globally by 2030 (Bloomberg 2018).

A service provider of EV charging faces significant operational challenges. First, there are stochastic arrivals of customers who have distinctive arrival time, desired departure time, and charging requirements for their EVs, observations which we show using real data in the numerical study section. Second, the tariff structure for an EV charging service provider in the U.S.—a commercial electricity consumer—includes *demand charges*, i.e., costs "based on the highest average electricity usage occurring within a defined time interval (usually 15 minutes) during a billing period" (National Renewable Energy Laboratory 2017). The total demand charge is the sum of that applicable to all periods during the entire horizon and that applicable to periods during only on-peak or mid-peak hours: Time-differentiated demand charges were developed to alleviate the issue that all-period demand charges may penalize users with peak load, which may not coincide with the system peak (Hausman and Neufeld 1984, Veall 1983). This total demand charge for an EV charging service provider can be as high as 70% of its total electricity cost (Chitkara et al. 2016).

In this paper, we study the problem of a service provider in scheduling EV charging with the goal of minimizing the total expected cost. We model it as a stochastic program (SP). We characterize the random number of arriving customers to follow Poisson distributions, commonly used for customer arrivals into a charging station (Mak et al. 2013, Schneider et al. 2017, He et al. 2021).

This SP is large-scale due to a large number of customer types which differ in arrival time, departure time, and charging quantity. Therefore, we propose to solve it approximately using a new approach called exponential cone programming (Chares 2009). It belongs to the class of conic programming (Boyd and Vandenberghe 2004) that generalizes linear programming to incorporate inequalities defined by an exponential cone—a three-dimensional convex cone involving exponentials and logarithms. Exponential cone programming is an attractive approach as it can be solved efficiently (in polynomial time) owing to the recent availability of solvers, such as MOSEK (MOSEK ApS 2020). There have been only a few attempts to use exponential cone programming in real applications, such as inventory management, manpower planning, and insurance reimbursement (See and Sim 2010, Jaillet et al. 2021, Zhu et al. 2021). Among these, only See and Sim (2010) involves solving an SP, and they use an approach based on exponential cone programming without explicitly formulating an exponential cone program (ECP). In addition, See and Sim (2010) requires the random variables in the SP to be sub-Gaussian, different from our paper, and thus cannot be directly applied to our EV charging problem to derive ECPs. Therefore, in this paper, we develop new methods to derive our ECPs to solve our SP approximately. In particular, we utilize moment-generating functions (MGFs) of the random variables to bound an order statistic in the objective introduced by the demand charge.

We then benchmark our ECP approach to solve the SP with two common approaches, sample average approximation (SAA) and a distributionally robust optimization (DRO) approach using a semi-definite program (SDP). We compare these approaches on numerical instances calibrated to real data of electricity tariff and customer arrivals. We summarize our theoretical and numerical contributions as follows:

- We derive an ECP approximation for the case in which the provider has unlimited chargers labeled ECP-U (where U represents "uncapacitated"). We show that the optimal expected cost of ECP-U gives an upper bound of that of our uncapacitated SP. We also characterize a bound on the gap between the theoretical performances of ECP-U and the SP.
- We derive another ECP approximation for the case with finite chargers, labeled ECP-C (where C represents "capacitated"), by combining the method in deriving ECP-U and also the idea from DRO of employing an entropic dominance ambiguity set (Chen et al. 2019). Instead of using DRO to mitigate distributional ambiguity, we use it to derive ECP-C as a tractable upper bound of the capacitated SP. We show that the optimal expected cost of ECP-C converges to a finite value bounded by that of ECP-U as the number of chargers tends to infinity.
- We extend our ECP approach to include also the pricing decision (together with the charging scheduling decisions) for both the uncapacitated and capacitated cases. We show that our ECP approximations can be solved efficiently using an alternating optimization algorithm.
- As our numerical instances are large-scale (with about 80,000 random variables and 700,000 decision variables), we show that while DRO-based SDP cannot be solved, ECP-C runs efficiently (about 50 times faster than SAA), and results in not only a lower standard deviation of the total cost but also a lower mean of this cost than SAA because SAA cannot be solved to optimality within a reasonable amount of time. We also find the optimality gap of ECP-C to be at most around 4% on all instances. In addition, we show the superior performance of ECP over SAA and other benchmarks continues to hold considering three practical implementation issues: finer time discretization, a data-driven setting where there are estimation errors in arrival rates or model mis-specification of arrival distribution, and an adaptive EV charging

environment. In the extended model with pricing decision, we show that our ECP approach continues to perform better than SAA.

• Using the ECP approach on our numerical instances also results in the following managerial insights: First, it is beneficial to incorporate customer departure information in scheduling optimization. Second, either an on-peak or mid-peak demand charge alone is not sufficient to smooth the electricity load, but an all-period demand charge alone can smooth the electricity load as well as when all three types of demand charge exist.

Our paper has the following academic and practice relevance: First, though developed for the EV charging problem, our method of constructing ECPs can be potentially applicable to approximate more general two-stage linear SPs with fixed recourse. Second, our ECP-C approximation can be used by charging service providers in managing EV charging due to its superior performance. Third, our numerical results provide insights for charging service providers on the value of customer departure information as well as for regulatory bodies on designing tariff structures in smoothing electricity load.

The rest of the paper is organized as follows. After reviewing the literature in §2, we introduce our SP in §3 for scheduling EV charging. In §4, we first develop ECP-U for the case in which the charging service provider has unlimited chargers, and provide its performance guarantee; we then develop ECP-C for the capacitated case. In §5, we consider two benchmark approaches to solving our SP model. In §6, we consider an extension of our model to jointly optimize the pricing and scheduling decisions and derive ECP approximations, which we solve by developing an alternating optimization algorithm. In §7 using numerical instances calibrated to real data, we compare the performance of ECP-C with other approaches and generate managerial insights using ECP-C. We conclude and discuss future work in §8. Any proof not in the main text can be found in the Online Appendix.

## 2. Literature Review

Our paper contributes to the literature on optimization with uncertainty. Common approaches include stochastic programming (Danzig 1955, Shapiro et al. 2009, Birge and Louveaux 2011), robust optimization (e.g., Ben-Tal et al. 2009, Goldfarb and Iyengar 2003, Bertsimas and Sim 2004, Ben-Tal et al. 2013) and, more recently, DRO (e.g., Delage and Ye 2010, Wiesemann et al. 2014, Esfahani and Kuhn 2018, Bertsimas et al. 2019, and the references therein). We build an SP model for the EV charging scheduling problem and specifically contribute to the literature of using the expected recourse function approximation in solving SPs (see Birge and Louveaux 2011, Chapter 8 for example) by proposing a new approach based on ECPs. ECPs have been used in a limited number of papers, such as See and Sim (2010), Jaillet et al. (2021), and Zhu et al. (2021), but none of the existing methods to derive ECPs applies to our EV charging problem. In particular, we develop ECP-U by using an MGF to bound the largest order statistic introduced by the demand charge. This method can be potentially applied to approximate more general two-stage linear SP with fixed recourse. We then combine this method together with the idea in the DRO framework of employing an entropic dominance ambiguity set (Chen et al. 2019) to derive ECP-C. In addition, all earlier work solves ECPs approximately using second-order cones or cutting plane due to a lack

of efficient solvers, while we solve ECPs exactly. Our paper also contributes to the nascent literature in EV operations management, which focuses on problems such as the planning and operations of battery swapping stations (Mak et al. 2013, Sun et al. 2019, Schneider et al. 2017), the adoption of EVs (Avci et al. 2014, Lim et al. 2014), charging infrastructure planning (He et al. 2021), and vehicle-to-grid operations (Zhang et al. 2021). In particular, He et al. (2021) present an integrated model to jointly determine the size and location of EV charging stations for a fleet of EVs in a vehicle sharing context. Zhang et al. (2021) formulate a two-stage stochastic integer program to study service-zone and facility capacity planning and fleet management in EV sharing systems with vehicle-to-grid integration. Different from those papers, which primarily focus on infrastructure planning, we optimize for a charging schedule that caters to various customer types differentiated by arrival/departure times and charging requirements.

In the literature on EV charging management from the perspective of a charging service provider, Jin et al. (2013) and Zhang et al. (2014) consider the problem of scheduling EV charging in the presence of an energy storage system. Our paper differs from the first in that our EV arrivals are stochastic while theirs is deterministic, and differs from the second in that we optimize the charging decisions while their charging rate is fixed. Jiang and Powell (2016) optimize the charging of one EV within a reservation window and consider the risk of the charging cost, while we model the setting where the provider charges multiple EVs with stochastic arrivals. Wu et al. (2021) study the problem of incentivizing customers to delay charging by designing a menu of price and departure time using a mechanism design approach, while we study the problem of optimizing the charging schedule with stochastic customer arrivals using a stochastic optimization approach. In addition, different from all these papers, we consider the setting where the provider buys electricity from a utility company with a tariff structure that includes demand charges.

Notation. We use boldface uppercase and lowercase characters to denote matrices and vectors, respectively. For example,  $x \in \mathbb{R}^n$  means that x is an *n*-dimensional real vector. For a matrix  $A \in \mathbb{R}^{m \times n}$ ,  $A'$  is the transpose. As usual, we use 1 to denote a vector of all 1's and 0 to denote the vector or a matrix of all 0's. We define the inner product  $X \cdot Y = \text{trace}(X'Y)$  for any  $X, Y \in \mathbb{R}^{m \times n}$ .

By  $x \geq y$ , we mean that x is greater than y component-wise. We denote by  $[N] \triangleq \{1, 2, ..., N\}$  the set of positive running indices up to *N*. We use  $\lvert \cdot \rvert$  ( $\lvert \cdot \rvert$ ) as a floor (ceil) function that takes a real number as input and gives as output the greatest (smallest) integer less (greater) than or equal to the input. We use  $\mathcal{P}_0(\mathcal{Z})$  to represent the set of all probability distributions on the support set  $\mathcal{Z}$ . A random vector,  $\tilde{z}$ , is denoted with a tilde sign and we use  $\tilde{z} \sim \mathbb{P}$ ,  $\mathbb{P} \in \mathcal{P}_0(\mathbb{R}^{I_z})$  to define  $\tilde{z}$  as an  $I_z$ -dimensional random variable with distribution  $\mathbb{P}$ . We use  $\mathbb{E}_{\mathbb{P}}[\cdot]$  to signify the expectation with respect to P.

## 3. A Stochastic Programming Model for Scheduling EV Charging

We model the operations of a charging service provider to minimize the total expected cost in a finite horizon of *T* periods. The service provider has a limited number of chargers, denoted by *C*. We assume customers do not enter a station that does not have chargers available. This assumption is due to the fact that EVs (e.g., Tesla) are equipped with onboard or mobile navigation systems that can help route them only to stations with available chargers.

As the EV charging industry is evolving quickly, there are currently various business practices. We consider one common practice where the service provider requests from each customer both *desired departure time* and *quantity to charge* (i.e., charging requirement) to manage the charging more effectively. For example, the adaptive charging network framework developed at  $Caltech<sup>1</sup>$  and the corresponding commercialized charging management software developed by Powerflex (2021) allow drivers to provide their estimated departure time and energy request through a mobile app. Moreover, this setting echoes the literature in EV charging scheduling (e.g., Jin et al. 2013, Xu et al. 2016). Providing such departure time information is intuitively more valuable for the case of long-duration charging, such as a few hours in the workplace and residential charging, than with short-duration charging, such as 30 minutes in public charging stations at highway stops. However, our model is general to accommodate both short- and long-duration charging. Further, we show using data in §7.5.1 even if the duration is short (such as 35 minutes on average), the value of providing departure information for the scheduling optimization is still significant, and this value increases rapidly when the duration increases. This alludes to the value of service providers requesting such departure information even when the charging duration is short.

EV arrivals. To characterize the customer arrival profile in the aforementioned general business practice, we classify EV customers into *V* types according to the triple  $(s_v, \tau_v, u_v)$ , where  $s_v$  is the arrival time,  $\tau_v$  is the desired departure time  $(s_v \leq \tau_v)$ , and  $u_v$  is the charging requirement for customer type  $v \in [V]$ . This definition of customer types is also consistent with real data used in §7, where customers are heterogeneous in all three dimensions.

<sup>1</sup> https://ev.caltech.edu/info

Recent research on EV operations (e.g., Mak et al. 2013, Schneider et al. 2017, He et al. 2021) often characterize customer arrivals at a charging (or battery swapping) station by Poisson distributions, which is empirically verified in He et al. (2021). Similarly, we assume the arrivals of customer types  $v \in [V]$  with infinite chargers are independent Poisson random variables, and each has an arrival rate  $\lambda_v$ . However, as the charging capacity is finite, the arrivals of customer types *v* (denoted by  $\tilde{z}_v$ ) are truncated Poisson random variables<sup>2</sup>, and thus they become dependent in this case. Let  $V_t \triangleq \{v \in [V] \mid s_v \le t \le \tau_v\}$  denote the set of customer types at the station in period *t*. Let  $\tilde{\mathbf{z}} \triangleq (\tilde{z}_v)_{v \in [V]}$  denote the vector of these truncated Poisson random variables drawn from the joint distribution denoted by  $\mathbb{P}^C$  (where *C* represents "capacitated") with arrival rate  $\lambda \triangleq (\lambda_v)_{v \in [V]}$  and support *Z* defined as:

$$
\mathcal{Z} \triangleq \left\{ \mathbf{z} \geq \mathbf{0} \mid \sum_{v \in \mathcal{V}_t} z_v \leq C, \ \forall t \in [T] \right\}.
$$
 (1)

Decisions. The service provider decides on a menu-based schedule where the charging decision is adaptive to customer types. Specifically, the menu-based charging schedule is  $\mathbf{x} \triangleq (x_{v,t})_{v \in [V], t \in \mathcal{T}_v}$ , where  $x_{v,t}$  denotes in period  $t \in \mathcal{T}_v$  the charging speed for customer type *v* (i.e., the quantity of electricity used to charge the EV of this type in this period) and  $\mathcal{T}_v \triangleq \{s_v, \ldots, \tau_v\}$  is the set of periods within the charging window of customer type *v*. Our modeling of the charging speed optimization is motivated by the observation in our data that the average charging speeds for most EVs vary and are *not* at their maximums, even for chargers with charging power greater than 22 kW (see §7). The sequence of events in each period  $t \in [T]$  is as follows: At the beginning of each period *t*, the charging service provider observes the realization of customer arrival  $\tilde{z}_v$  with  $s_v = t$ , and charges  $x_{v,t}$  to EVs of type  $v \in V_t$  by looking up the menu-based charging schedule x. At the end of period *t*, customers with desired departure time equal to *t* (i.e,  $\tau_v = t$ ) finish their charging service and leave the charging station.<sup>3</sup> The feasible set of  $x$ , denoted by  $\mathcal{X}$ , is given as follows:

$$
\mathcal{X} \triangleq \left\{ \boldsymbol{x} \mid \begin{array}{c} \sum_{t \in \mathcal{T}_v} \eta x_{v,t} = u_v \quad \forall v \in [V] \\ 0 \le x_{v,t} \le K/\eta \quad \forall v \in [V], t \in \mathcal{T}_v \end{array} \right\},\
$$

where the first constraint indicates that customer type *v* needs to fulfill the charging requirement  $u_v$  before  $\tau_v$  and  $\eta \in (0,1]$  is the ratio of the quantity of electricity increased in the battery to the quantity of electricity used to charge the battery. The second constraint indicates that the charging speed is within the limit, where *K* (in energy unit/period) is the maximum amount of energy a battery can be charged in each period.<sup>4</sup> For a well-defined type  $v \in [V]$ , it thus satisfies

<sup>2</sup> For customers arriving at the same time, we allow any way of truncation.

<sup>&</sup>lt;sup>3</sup> Note all of our analytical results can be extended in a straightforward manner when we modify our model to represent a steady state situation, where type *v* with  $\tau_v < s_v$  represents customers departing in period  $T + \tau_v$ .

 $4$  For the case in which  $\eta$  and  $K$  differ for different customer types (e.g., with various states of charge), our theoretical results in §4 and §6 extend in a straightforward manner.

 $u_v \le K(\tau_v - s_v + 1)$ , as otherwise the charging requirement cannot be met even with charging at the maximum speed for the entire stay duration.

Total cost. The objective of the service provider is to minimize its expected total electricity cost, which includes demand charge as mentioned in §1. This demand charge depends on the maximum per-period electricity used within the horizon and is imposed on almost all commercial electricity users (such as charging stations) by their utility firms in the U.S. (Mills et al. 2008, National Renewable Energy Laboratory 2017). The rationale is for utility firms to recover the maintenance cost of standby equipment in the event of high load from users and also incentivize them to smooth the electricity load over time. Let  $d$  (in  $\frac{2}{\pi}$ /(energy unit/period)) denote the unit demand charge; let  $f_t(\mathbf{x}, \tilde{\mathbf{z}})$  denote the total electricity used to charge EVs of all customers in period t, i.e.,

$$
f_t(\boldsymbol{x}, \tilde{\boldsymbol{z}}) = \sum_{v \in \mathcal{V}_t} x_{v,t} \tilde{z}_v.
$$

Given  $(x, \tilde{z})$ , the total cost of charging all EVs within the horizon, denoted by  $c(x, \tilde{z})$ , is as follows:

$$
c(\boldsymbol{x}, \tilde{\boldsymbol{z}}) \triangleq d \max_{t \in [T]} \{ f_t(\boldsymbol{x}, \tilde{\boldsymbol{z}}) \} + \sum_{s \in [T]} e_s f_s(\boldsymbol{x}, \tilde{\boldsymbol{z}}), \tag{2}
$$

where the first term is the demand charge, and the second term is the energy charge, with *e<sup>t</sup>* (in \$/energy unit) denoting the unit time-of-use (TOU) energy charge. Note that "demand charge" and "energy charge" are terminology used in the rate structure of utility firms, where "demand" differs from the usual meaning in "customer demand," and "charge" differs from that used in "EV charging." Note also that for simplicity of exposition, the demand charge in (2) consists of only an all-period demand charge, applicable to all periods in the finite horizon. We consider the case in which the total demand charge is the sum of multiple types of demand charge, such as demand charge in peak hours and mid-peak hours, in the numerical study in §7. Lastly, note that our model and theoretical results can be easily extended to the case where demand charge depends on the maximum electricity usage *per multiple periods* (see an example in §7.4.1).

Model Formulation. We formulate the problem of scheduling EV charging as an SP:

$$
\min_{x \in \mathcal{X}} \mathbb{E}_{\mathbb{P}^C}[c(\boldsymbol{x}, \tilde{\boldsymbol{z}})],\tag{3}
$$

where the expectation is over  $\tilde{z}$ . We denote an optimal solution to (3) by  $x^*$  and the optimal value of (3) by  $\pi^*$ , i.e.,  $\pi^* = \mathbb{E}_{\mathbb{P}^C} [c(\bm{x}^*, \tilde{\bm{z}})].$ 

Note that  $c(\mathbf{x}, \tilde{\mathbf{z}})$  is a nonlinear function which can be linearized with auxiliary variables, so problem (3) can be rewritten as a two-stage stochastic linear program with complete recourse. Two-stage SPs are in general #P-hard, as evaluating the expectation of the random objective function is #P-hard (Hanasusanto et al. 2016). In addition, our SP model (3) is large-scale due to its large number of random variables, *V* , and decision variables: *V* is the product of three terms, the number of possible arrival periods, departure periods (given an arrival period), and charging requirements; the latter is the product of *V* and the number of periods within the charging window of all types. As we show in our numerical study calibrated to real data, for a representative one-day planning horizon, the numbers of random variables and decision variables are about 80,000 and 700*,* 000, respectively. We next propose an approach to solve our SP based on ECP approximation in §4, and benchmark ECP with two common approaches, SAA and DRO, which we discuss in §5.

## 4. Exponential Cone Programming Approximations

In this section, we first consider the uncapacitated case by developing an ECP approximation and providing a theoretical performance guarantee. We then extend it to the capacitated case by using the idea from the framework of DRO.

As mentioned before, an ECP is a conic program with constraints that can be represented by exponential cones, which are three-dimensional convex cones involving exponentials and logarithms:

$$
\mathcal{K}_{\exp} \triangleq \left\{ (x_1, x_2, x_3) \mid x_1 \geq x_2 \exp(x_3/x_2), x_2 > 0 \right\} \cup \left\{ (x_1, 0, x_3) \mid x_1 \geq 0, x_3 \leq 0 \right\}.
$$

ECP generalizes the widely used LP and second-order conic program (SOCP) (see, e.g., Chandrasekaran and Shah 2017). ECPs can be solved in polynomial time using an interior point algorithm with an efficient solver, such as MOSEK. Note that since the most difficult step in solving (3) is computing the expectation of the largest order statistic introduced by the demand charge (while the energy charge is linear in  $\tilde{z}$ ), ECPs become a natural choice for the approximation of (3) because we use the MGF (involving exponential functions) of  $\tilde{z}$  to bound this order statistic.

#### 4.1. The Uncapacitated Case

Let  $\mathbb{P}^{\infty}$  denote the joint distribution of  $\tilde{z}$  for  $C = \infty$ . Note that  $\tilde{z}_v$ 's for  $v \in [V]$  are independent Poisson random variables. Given any  $\boldsymbol{\theta} \triangleq (\theta_v)_{v \in [V]}$ , the MGF of  $\tilde{z} \sim \mathbb{P}^{\infty}$  is

$$
\mathbb{E}_{\mathbb{P}^{\infty}}\left[\exp\left(\sum_{v\in[V]} \theta_v \tilde{z}_v\right)\right] = \prod_{v\in[V]} \mathbb{E}_{\mathbb{P}^{\infty}}\left[\exp\left(\theta_v \tilde{z}_v\right)\right] = \prod_{v\in[V]} \exp\left(\lambda_v(e^{\theta_v}-1)\right),\tag{4}
$$

where the first equality is due to the independence of  $\tilde{z}_v$ 's and the second equality follows from the closed-form MGF expression of a Poisson random variable  $\tilde{z} \sim \mathbb{P}$  with arrival rate  $\lambda$ , i.e.,  $\mathbb{E}_{\mathbb{P}}[e^{\theta \tilde{z}}] = e^{\lambda(e^{\theta}-1)}$  for any  $\theta$ . Since (4) involves exponential functions, we obtain an upper bound of  $\pi^*$ , the optimal value of (3), using an ECP as follows:

PROPOSITION 1. *When*  $C = \infty$ , the optimal value of ECP-U gives an upper bound of  $\pi^*$ :

$$
\inf_{\mathbf{x}\in\mathcal{X},\kappa,\gamma,\mu>0,\xi,\zeta} d(\kappa+\gamma)+\sum_{s\in[T]} e_s f_s(\mathbf{x},\boldsymbol{\lambda})
$$
\n
$$
\text{s.t.} \sum_{v\in\mathcal{V}_t} x_{v,t}\lambda_v \le \gamma \qquad \qquad \forall t\in[T] \tag{5a}
$$

$$
(ECP-U) \qquad \mu \exp(x_{v,t}/\mu) \le \xi_{v,t} \qquad \forall t \in [T], v \in \mathcal{V}_t \qquad (5b)
$$

$$
\mu \exp\left(\left(-\kappa + \sum_{v \in \mathcal{V}_t} \lambda_v(\xi_{v,t} - x_{v,t} - \mu)\right) / \mu\right) \le \zeta_t \qquad \forall t \in [T] \qquad (5c)
$$

$$
\sum_{t \in [T]} \zeta_t \le \mu \tag{5d}
$$

*Proof.* To obtain an upper bound of  $\pi^*$ , we first obtain an upper bound of  $\mathbb{E}_{\mathbb{P}^\infty}\left[\max_{t\in[T]} f_t(\bm{x}, \tilde{\bm{z}})\right]$ :

$$
\mathbb{E}_{\mathbb{P}^{\infty}}\left[\max_{t\in[T]}f_t(\boldsymbol{x},\tilde{\boldsymbol{z}})\right] = \mathbb{E}_{\mathbb{P}^{\infty}}\left[\max_{t\in[T]} \left(f_t(\boldsymbol{x},\tilde{\boldsymbol{z}})-f_t(\boldsymbol{x},\boldsymbol{\lambda})+f_t(\boldsymbol{x},\boldsymbol{\lambda})\right)\right] \n\leq \mathbb{E}_{\mathbb{P}^{\infty}}\left[\max_{t\in[T]} \left(f_t(\boldsymbol{x},\tilde{\boldsymbol{z}})-f_t(\boldsymbol{x},\boldsymbol{\lambda})\right)+\max_{t\in[T]} f_t(\boldsymbol{x},\boldsymbol{\lambda})\right] \n= \mathbb{E}_{\mathbb{P}^{\infty}}\left[\max_{t\in[T]} \left(f_t(\boldsymbol{x},\tilde{\boldsymbol{z}})-f_t(\boldsymbol{x},\boldsymbol{\lambda})\right)\right]+\max_{t\in[T]} f_t(\boldsymbol{x},\boldsymbol{\lambda}).
$$
\n(6)

We then obtain an upper bound of the first term in (6) given any  $\mu > 0$  as follows:

$$
\mathbb{E}_{\mathbb{P}^{\infty}}\left[\max_{t\in[T]}(f_t(\boldsymbol{x}, \tilde{\boldsymbol{z}}) - f_t(\boldsymbol{x}, \boldsymbol{\lambda}))\right] \leq \mu \ln \mathbb{E}_{\mathbb{P}^{\infty}}\left[\exp\left(\max_{t\in[T]}(f_t(\boldsymbol{x}, \tilde{\boldsymbol{z}}) - f_t(\boldsymbol{x}, \boldsymbol{\lambda}))/\mu\right)\right]
$$
  
\n
$$
\leq \mu \ln \mathbb{E}_{\mathbb{P}^{\infty}}\left[\sum_{t\in[T]} \exp\left((f_t(\boldsymbol{x}, \tilde{\boldsymbol{z}}) - f_t(\boldsymbol{x}, \boldsymbol{\lambda}))/\mu\right)\right]
$$
  
\n
$$
= \mu \ln \sum_{t\in[T]} \mathbb{E}_{\mathbb{P}^{\infty}}\left[\exp\left((f_t(\boldsymbol{x}, \tilde{\boldsymbol{z}}) - f_t(\boldsymbol{x}, \boldsymbol{\lambda}))/\mu\right)\right]
$$
  
\n
$$
= \mu \ln \sum_{t\in[T]} \mathbb{E}_{\mathbb{P}^{\infty}}\left[\exp\left(\sum_{v\in\mathcal{V}_t} \frac{x_{v,t}}{\mu} (\tilde{\boldsymbol{z}}_v - \lambda_v)\right)\right]
$$
  
\n
$$
= \mu \ln \sum_{t\in[T]} \prod_{v\in\mathcal{V}_t} \mathbb{E}_{\mathbb{P}^{\infty}}\left[\exp\left(\frac{x_{v,t}}{\mu} \tilde{\boldsymbol{z}}_v\right)\right] \exp\left(-\frac{x_{v,t}}{\mu} \lambda_v\right)
$$
  
\n
$$
= \mu \ln \sum_{t\in[T]} \exp\left(\sum_{v\in\mathcal{V}_t} \lambda_v \left(e^{x_{v,t}/\mu} - 1 - x_{v,t}/\mu\right)\right), \tag{7}
$$

where the first line is due to the convexity of an exponential function and Jensen's inequality:

$$
\exp\left(\mathbb{E}_{\mathbb{P}^\infty}\left[\max_{t\in[T]}(f_t(\boldsymbol{x}, \tilde{\boldsymbol{z}})-f_t(\boldsymbol{x}, \boldsymbol{\lambda})/\mu)\right]\right)\leq \mathbb{E}_{\mathbb{P}^\infty}\left[\exp\left(\max_{t\in[T]}(f_t(\boldsymbol{x}, \tilde{\boldsymbol{z}})-f_t(\boldsymbol{x}, \boldsymbol{\lambda})/\mu)\right)\right];
$$

the fourth line results from substituting  $f_t(\mathbf{x}, \tilde{\mathbf{z}}) = \sum_{v \in \mathcal{V}_t} x_{v,t} \tilde{z}_v$  into the right-hand side (RHS) of the third line; and the fifth and sixth lines follow from (4).

Note if we combine (6) and (7) as follows

$$
\mathbb{E}_{\mathbb{P}^\infty}\left[\max_{t\in[T]}f_t(\boldsymbol{x}, \tilde{\boldsymbol{z}})\right] \leq \inf_{\mu>0}\left(\mu\ln\sum_{t\in[T]}\exp\left(\sum_{v\in\mathcal{V}_t}\lambda_v\left(e^{x_{v,t}/\mu}-1-x_{v,t}/\mu\right)\right)\right) + \max_{t\in[T]}\tf_t(\boldsymbol{x}, \boldsymbol{\lambda}),
$$

its RHS has the following epigraph form

$$
\inf_{\mu>0,\kappa,\gamma} \kappa + \gamma
$$
\n
$$
\text{s.t.} \sum_{v \in \mathcal{V}_t} x_{v,t} \lambda_v \le \gamma, \quad \forall t \in [T]
$$
\n
$$
\mu \ln \sum_{t \in [T]} \exp \left( \sum_{v \in \mathcal{V}_t} \lambda_v \left( e^{x_{v,t}/\mu} - 1 - x_{v,t}/\mu \right) \right) \le \kappa. \tag{8}
$$

The constraint (8) is equivalent to  $\sum_{t \in [T]} \mu \exp \left( \sum_{v \in \mathcal{V}_t} \lambda_v \left( e^{x_{v,t}/\mu} - x_{v,t}/\mu - 1 \right) - \kappa/\mu \right) \le$  $\mu$ , where the expression within the first summation of the left-hand side, i.e.,  $\mu \exp\left(\sum_{v \in \mathcal{V}_t} \lambda_v \left(e^{x_{v,t}/\mu} - x_{v,t}/\mu - 1\right) - \kappa/\mu\right)$ , has an epigraph form for all  $t \in [T]$  as follows:

$$
\inf_{\zeta_t} \quad \zeta_t
$$
\n
$$
\text{s.t. } \mu \exp\left(\left(-\kappa + \sum_{v \in \mathcal{V}_t} \lambda_v(\mu e^{x_{v,t}/\mu} - x_{v,t} - \mu)\right) / \mu\right) \le \zeta_t. \tag{9}
$$

Further, the term  $\mu e^{x_{v,t}/\mu}$  inside constraint (9) has an epigraph form  $\inf_{\xi_{v,t}} \left\{ \xi_{v,t} : \mu e^{x_{v,t}/\mu} \le \xi_{v,t} \right\}$ for all  $v \in V_t$ . Hence, the epigraph containing constraint (8) can be equivalently written as

$$
\inf_{\mu>0,\kappa,\gamma,\xi,\zeta} \kappa + \gamma
$$
  
s.t. (5a) – (5d). (10)

Since  $f_s(\mathbf{x}, \tilde{\mathbf{z}})$  is linear in  $\tilde{\mathbf{z}}$  for all  $s \in [T]$  and  $\tilde{\mathbf{z}} \sim \mathbb{P}^\infty$  is the vector of independent Poisson random variables, we have

$$
\mathbb{E}_{\mathbb{P}^{\infty}}\left[f_s(\boldsymbol{x}, \tilde{\boldsymbol{z}})\right] = f_s(\boldsymbol{x}, \boldsymbol{\lambda}).\tag{11}
$$

Therefore, by combining  $(6)$ ,  $(7)$ ,  $(10)$ , and  $(11)$  and optimizing over  $x$ , we get the optimal value of ECP-U as an upper bound of  $\pi^*$ , as a solution optimal to ECP-U is also feasible to (3).  $\Box$ 

Note that the model in Proposition 1 is labeled ECP-U (recall U represents "uncapacitated") because all the constraints in this model involve either linear or exponential functions and thus can be expressed as exponential cone constraints. For example, the constraint (5b),  $\mu \exp(x_{v,t}/\mu) \leq \xi_{v,t}$ , can be written as  $(\xi_{v,t}, \mu, x_{v,t}) \in \mathcal{K}_{\text{exp}}$ . Hence, ECP-U is an ECP, and thus can be solved via a stateof-the-art conic programming solver efficiently (MOSEK ApS 2020). Note also that the number of decision variables and constraints of ECP-U are on the same order as those of our SP model (3).

REMARK 1. To the best of our knowledge, the only other paper to approximate an SP with ECPs is See and Sim (2010), but their approach requires the random variables of the SP to be sub-Gaussian in order to derive meaningful bounds, which is not applicable to our SP. Moreover, our approximation method differs from theirs by directly bounding the expectation of piecewise linear convex functions while they first use a linear decision rule and then approximate the resulting expected surplus of affine functions. Finally, we directly solve ECPs, while all previous papers (e.g., See and Sim 2010) approximate ECPs using second-order cone programs.

REMARK 2. Though developed for a specific charging problem, our ECP approach of bounding the expected recourse function could be potentially useful for solving more general two-stage SPs. Consider a two-stage linear SP with fixed recourse where the recourse function is defined generally as

$$
Q(\boldsymbol{x}, \tilde{\boldsymbol{z}}) = \inf_{\boldsymbol{y}} \left\{ d'\boldsymbol{y} : \boldsymbol{H}(\tilde{\boldsymbol{z}}) \boldsymbol{x} + \boldsymbol{W} \boldsymbol{y} \geq \boldsymbol{h}(\tilde{\boldsymbol{z}}) \right\},
$$

where  $h(\tilde{z})$  and  $H(\tilde{z})$  are affine in  $\tilde{z}$ . If the dual feasible set  $\{q \ge 0 : W'q = d\}$  is non-empty and compact with vertices  $v_i$ ,  $i \in [I]$ , the expected recourse function becomes

$$
\mathbb{E}_{\mathbb{P}}\left[Q(\boldsymbol{x}, \tilde{\boldsymbol{z}})\right] = \mathbb{E}_{\mathbb{P}}\left[\max_{i \in [I]} \boldsymbol{v}_i' \left(\boldsymbol{h}(\tilde{\boldsymbol{z}}) - \boldsymbol{H}(\tilde{\boldsymbol{z}})\boldsymbol{x}\right)\right].
$$

If  $\tilde{z}$  has independent components with exponential cone representable log-MGF, this expected recourse function can be upper bounded by an ECP using the same approach as in Proposition 1.

We denote an optimal solution to ECP-U by  $\bar{x}$  and the value of the objective function in (3) evaluated at  $\bar{x}$  by  $\pi^E$ , i.e.,  $\pi^E = \mathbb{E}_{\mathbb{P}^\infty} [c(\bar{x}, \tilde{z})]$ . We next show that compared with an optimal solution to (3),  $x^*$ , there is a theoretical guarantee on the performance of  $\bar{x}$ . Theorem 1 compares  $\pi^E$  and  $\pi^*$  (note that  $\pi^* = \mathbb{E}_{\mathbb{P}^\infty} [c(\mathbf{x}^*, \tilde{\mathbf{z}})] = \mathbb{E}_{\mathbb{P}^C} [c(\mathbf{x}^*, \tilde{\mathbf{z}})]$ , as  $\mathbb{P}^C \equiv \mathbb{P}^\infty$  when *C* is infinite).

THEOREM 1. *Any optimal solution*  $\bar{x}$  to *ECP-U has a performance quarantee as follows:* 

$$
\pi^{E} - \pi^{*} = \mathbb{E}_{\mathbb{P}^{\infty}} \left[ c(\bar{x}, \tilde{z}) \right] - \mathbb{E}_{\mathbb{P}^{\infty}} \left[ c(\bm{x}^{*}, \tilde{z}) \right] \le d \inf_{\mu > 0} \left( \mu \ln \sum_{t \in [T]} \exp \left( \sum_{v \in \mathcal{V}_{t}} \lambda_{v} \phi \left( \frac{\min\{u_{v}, K\}}{\mu \eta} \right) \right) \right), \quad (12)
$$

*where*  $\phi(x) \triangleq e^x - 1 - x$ .

Theorem 1 gives in (12) an upper bound on the difference between  $\pi^{E}$  and  $\pi^{*}$ . This bound depends on *d*,  $\eta$ , *K*, and *T*, as well as  $\lambda_v$  and  $u_v$  (for all  $v \in [V]$ ), but it does not depend on  $e_s$ . Although this bound is not tight, it can be computed *before* solving ECP-U. *After* solving ECP-U, the bound on the difference between  $\pi^E$  and  $\pi^*$  on our numerical instances becomes tighter (see §7).

#### 4.2. The Capacitated Case

Unlike the uncapacitated case,  $\tilde{z} \sim \mathbb{P}^C$  does not follow independent Poisson distributions when  $C < \infty$ , so it is difficult to obtain its analytical characterization analogous to that in (4). We then use the idea of DRO: It is often used to mitigate distributional ambiguity, but we use DRO to derive an ECP approximation as a tractable upper bound of our SP model. Note that in our SP model the arrival distribution is *known* (joint truncated Poisson distribution,  $\tilde{z} \sim \mathbb{P}^C$ ), and thus there is no distributional ambiguity. We relax this known distribution to be contained in an ambiguity set in the DRO framework, which we denote by  $\mathcal{F}$ , i.e.,  $\mathbb{P}^C \in \mathcal{F}$ . Therefore, the solution from the DRO framework gives an upper bound of our SP model with the known distribution. As we show below, the DRO formulation is intractable, and thus we obtain an upper bound of this DRO formulation by developing an ECP approximation, which in turn gives an upper bound of our SP model.

We use the infinitely constrained "entropic dominance" ambiguity set, adapted from Chen et al. (2019):

$$
\mathcal{F} \triangleq \left\{ \mathbb{P} \in \mathcal{P}_0 \left( \mathbb{R}^V \right) \middle| \begin{array}{c} \tilde{z} \sim \mathbb{P} \\ \ln \mathbb{E}_{\mathbb{P}} \left[ \exp \left( \boldsymbol{\theta}' \tilde{z} \right) \right] \leq \sum_{v \in [V]} \lambda_v \left( e^{\theta_v} - 1 \right), \forall \boldsymbol{\theta} \geq \mathbf{0} \\ \mathbb{P} \left[ \tilde{z} \in \mathcal{Z} \right] = 1 \end{array} \right\},
$$

where the third line is the support constraint with the support  $Z$  given in (1). The second line is entropic dominance constraints, which restrict *all* the moments of random variables  $z_v$ 's to be no greater than those of independent Poisson variables (i.e., in uncapacitated case): The left- and right-hand sides of this set of constraints are the log-MGFs of  $\tilde{z} \sim \mathbb{P}^C$  and  $\tilde{z} \sim \mathbb{P}^{\infty}$  (see (4)), respectively. To further understand  $F$ , we write it as the intersection of the following two sets, as shown in Lemma 1:

$$
\mathcal{F}^1 \triangleq \left\{\mathbb{P} \in \mathcal{P}_0\left(\mathbb{R}^V\right) \middle| \; \tilde{\boldsymbol{z}} \sim \mathbb{P}; \; \ln \mathbb{E}_{\mathbb{P}}\left[\exp\left(\boldsymbol{\theta}^{\prime} \tilde{\boldsymbol{z}}\right)\right] \leq \sum_{v \in [V]} \lambda_v \left(e^{\theta_v} - 1\right), \forall \boldsymbol{\theta} \geq \mathbf{0} \right\}
$$
\n
$$
\mathcal{F}^2 \triangleq \left\{\mathbb{P} \in \mathcal{P}_0\left(\mathbb{R}^V\right) \middle| \; \tilde{\boldsymbol{z}} \sim \mathbb{P}; \; \mathbb{E}_{\mathbb{P}}\left[\tilde{\boldsymbol{z}}\right] \leq \boldsymbol{\lambda}; \; \mathbb{P}\left[\tilde{\boldsymbol{z}} \in \mathcal{Z}\right] = 1\right\}.
$$

LEMMA 1.  $\mathcal{F} = \mathcal{F}^1 \cap \mathcal{F}^2$ .

Lemma 1 shows the following: (1) Both the entropic dominance constraints (in  $\mathcal{F}^1$ ) and the support constraint (in  $\mathcal{F}^2$ ) are crucial in defining the ambiguity set. For instance, we show numerically that a model without  $\mathcal{F}^1$  performs much worse in §7.4.4. (2) As  $\mathcal{F}^2$  is an example of the moment-based ambiguity sets commonly used in DRO which involves only a constraint on the first moment (i.e.,  $\mathbb{E}_{\mathbb{P}}[\tilde{z}] \leq \lambda$ ), Lemma 1 implies that *F*, as a subset of  $\mathcal{F}^2$ , incorporates more moment information than  $\mathcal{F}^2$  does. Therefore,  $\mathcal F$  is smaller and less conservative than  $\mathcal{F}^2$ . Next, we show that the ambiguity set  $\mathcal F$  contains the given distribution  $\mathbb P^C$ .

PROPOSITION 2.  $\mathbb{P}^C \in \mathcal{F}$ .

Therefore, we can obtain an upper bound of the optimal value of (3) by considering the worst-case expected total cost over the ambiguity set *F*:

$$
\min_{\boldsymbol{x}\in\mathcal{X}}\sup_{\mathbb{P}\in\mathcal{F}}\mathbb{E}_{\mathbb{P}}\left[c(\boldsymbol{x},\tilde{\boldsymbol{z}})\right],\tag{DRO-Ent}
$$

where "Ent" represents "entropic." However, due to the infinite number of constraints in  $\mathcal F$  (since the entropic dominance constraints have to hold for all  $\theta \ge 0$ ), the DRO-Ent model is not tractable (Chen et al. 2019). Therefore, we approximate DRO-Ent using an ECP formulation similar to that in Proposition 1, which can also be solved efficiently. We use Lemma 1 to write DRO-Ent as the optimization over ambiguity set  $\mathcal{F}^1$  and  $\mathcal{F}^2$  separately. For the former case, we can use a technique similar to that of the proof of Proposition 1; for the latter, as  $\mathcal{F}^2$  is an example of moment-based ambiguity sets used in DRO, we use techniques standard in DRO to obtain bounds. We combine these two to obtain a tractable upper bound as follows, which we label as ECP-C (where C represents "capacitated" as mentioned above).

PROPOSITION 3. When  $C < \infty$ , the following ECP-C gives an upper bound of  $\pi^*$ :

 $v \in \mathcal{V}_t$ 

$$
\inf_{\substack{x \in \mathcal{X}, a, b \ge 0, \nu \ge 0, \mu > 0, \\ \kappa, \gamma, \alpha, \beta \ge 0, \xi, \zeta, \rho \ge 0}} \frac{\inf}{\kappa, \gamma, \alpha, \beta \ge 0, \mu > 0, d} (\kappa + \gamma + \alpha + \beta' \lambda) + a + b' \lambda
$$
\n
$$
\text{s.t.} \sum y_{v,t} \lambda_v \le \gamma \qquad \forall t \in [T] \tag{13a}
$$

$$
(\xi_{v,t}, \mu, y_{v,t}) \in \mathcal{K}_{\text{exp}} \qquad \qquad \forall t \in [T], v \in \mathcal{V}_t \qquad (13b)
$$

$$
\left(\zeta_t, \mu, -\kappa + \sum_{v \in \mathcal{V}_t} \lambda_v(\xi_{v,t} - y_{v,t} - \mu)\right) \in \mathcal{K}_{\exp}
$$
  $\forall t \in [T]$  (13c)

$$
(ECP-C) \qquad \sum_{t \in [T]} \zeta_t \le \mu \tag{13d}
$$

$$
C\sum_{k\in[T]}\rho_t^k \le \alpha \qquad \qquad \forall t\in[T] \qquad (13e)
$$

$$
x_{v,t} - y_{v,t} - \beta_v \le \sum_{k \in \mathcal{T}_v} \rho_t^k \qquad \qquad \forall t \in [T], v \in \mathcal{V}_t \qquad (13f)
$$

$$
C\sum_{t\in[T]} \nu_t \le a\tag{13g}
$$

$$
\sum_{s \in \mathcal{T}_v} x_{v,s} e_s - b_v \le \sum_{t \in \mathcal{T}_v} \nu_t \qquad \forall v \in [V] \qquad (13h)
$$

Similar to ECP-U, ECP-C can be solved efficiently. (Note that the number of decision variables and constraints in ECP-C is on the same order as those in our SP model (3).)

Theorem 2. *The optimal value of ECP-C increases as C increases and converges to a value that is upper bounded by the optimal value of ECP-U as*  $C \rightarrow \infty$ .

Theorem 2 shows that the optimal value of ECP-C increases as charging station capacity increases, intuitively because more vehicles will be admitted to the station to be charged. It also shows that this optimal value of ECP-C converges and yields an asymptotic lower bound of the optimal value of ECP-U. However, unlike ECP-U, we are unable to obtain a theoretical performance guarantee of an optimal action to ECP-C and thus use a numerical study to examine its performance together with those of two benchmark approaches.

## 5. Benchmark Approximations

## 5.1. Sample Average Approximation (SAA)

The most common approach to solve a two-stage SP is SAA based on Monte Carlo sampling: Generate from the given distribution many independent samples and replace this distribution in the SP model with the empirical distribution of these samples. We denote these samples from  $\mathbb{P}^C$  by  $\mathbf{z}^i \equiv (z_v^i)_{v \in [V]}$   $\forall i \in [N]$ , where *N* is the sample size. The SAA model can be obtained after reformulating (3) as a two-stage SP using auxiliary variables:

$$
\inf_{\mathbf{x}\in\mathcal{X},\gamma^{i}\forall i\in[N]} \frac{1}{N} \sum_{i\in[N]} \left( d\gamma^{i} + \sum_{s\in[T]} e_s \left( \sum_{v\in\mathcal{V}_s} x_{v,s} z_v^{i} \right) \right)
$$
\n
$$
\text{s.t.} \quad \sum_{v\in\mathcal{V}_t} x_{v,t} z_v^{i} \leq \gamma^{i} \qquad \forall t\in[T], i\in[N] \tag{14}
$$

Let us denote by  $x^S$  an optimal solution to (14) and by  $\pi^S$  the value of the objective function in (3) evaluated at  $x^S$ , i.e.,  $\pi^S = \mathbb{E}_{\mathbb{P}^C} [c(x^S, \tilde{z})]$ . Since  $x^S$  and  $\pi^S$  depend on the samples generated and are thus random, so to obtain a solution near optimal to (3) with a high probability requires the sample size N to be sufficiently large, especially given the large-scale nature of  $(3)$ . In particular, as mentioned in Shapiro and Nemirovski (2005), the sample complexity of obtaining an  $\epsilon$ -optimal solution is  $\tilde{O}(\frac{D^2 L^2}{\epsilon^2})$ , where *L* is the Lipchitz constant of the objective function  $f(\cdot, z)$  for all  $z \in \mathcal{Z}$ and  $D$  is the diameter of  $\mathcal X$ . Nonetheless, fewer samples may be needed in practice. For instance, in our numerical study, we need about 38,000 samples to obtain a near-optimal solution (see Figure 3 in §7.3).

#### 5.2. Distributionally Robust Optimization (DRO)

Another benchmark is a DRO approach with the mean-variance ambiguity set:

$$
\min_{\boldsymbol{x}\in\mathcal{X}}\sup_{\mathbb{P}\in\mathcal{F}_{\Sigma}}\mathbb{E}_{\mathbb{P}}\left[c(\boldsymbol{x},\tilde{\boldsymbol{z}})\right],\tag{DRO-C}
$$

where

$$
\mathcal{F}_{\Sigma} = \left\{ \mathbb{P} \in \mathcal{P}_{0} \left( \mathbb{R}^{V} \right) \middle| \begin{array}{c} \tilde{z} \sim \mathbb{P} \\ \mathbb{E}_{\mathbb{P}} \left[ \tilde{z} \right] = \mu \\ \mathbb{E}_{\mathbb{P}} \left[ \tilde{z} \tilde{z}' \right] \preceq \Sigma + \mu \mu' \\ \mathbb{P} \left[ \tilde{z} \in \mathcal{Z} \right] = 1 \end{array} \right\} . \tag{15}
$$

Using standard techniques in DRO (see Wiesemann et al. 2014, for example), the model DRO-C can be reformulated as the following SDP (see Online Appendix A for its derivation):

$$
\min_{\substack{\mathbf{x} \in \mathcal{X}, \alpha, \beta, \Gamma, \\ \mathbf{t}^t \geq \mathbf{0}, \mathbf{\zeta}^t \geq \mathbf{0}, \mathbf{y}^t, \forall t \in [T]}} d(\alpha + \beta' \boldsymbol{\mu} + \boldsymbol{\Gamma} \bullet (\boldsymbol{\mu} \boldsymbol{\mu}' + \boldsymbol{\Sigma})) + \sum_{s \in [T]} e_s f_s(\mathbf{x}, \boldsymbol{\mu})
$$
\n
$$
\xi^t \geq \mathbf{0}, \mathbf{\zeta}^t \geq \mathbf{0}, \mathbf{y}^t, \forall t \in [T]
$$
\n
$$
\text{s.t.} \qquad \begin{pmatrix} \alpha - C \sum_{s \in [T]} \zeta_s^t \ \mathbf{y}^{t'} \\ \mathbf{y}^t \end{pmatrix} \succeq \mathbf{0} \qquad \forall t \in [T]
$$
\n
$$
2y_v^t = \beta_v - x_{v,t} - \xi_v^t + \sum_{s \in \mathcal{T}_v} \zeta_s^t \qquad \forall t \in [T], \forall v \in \mathcal{V}_t
$$
\n
$$
2y_v^t = \beta_v - \xi_v^t + \sum_{s \in \mathcal{T}_v} \zeta_s^t \qquad \forall t \in [T], \forall v \in [V] \setminus \mathcal{V}_t.
$$

Note the above SDP formulation has *T* semidefinite constraints of dimension  $V + 1$ , and thus may run into computation time and memory issues when *V* and *T* are large. Note also in general the complexity of SDP is higher than ECP (which is the same as SOCP): Obtaining  $\epsilon$ -optimal solutions to SDP can be found in  $O(n^2k^{2.5}\log(1/\epsilon))$  arithmetic operations via interior point methods, where *n* is the number of decision variables and *k* is the dimension of the symmetric matrices in the SDP constraints; obtaining  $\epsilon$ -optimal solutions to ECPs requires  $O(n^2k^{1.5}\log(1/\epsilon))$  arithmetic operations, where *n* is the number of decision variables and *k* is the number of exponential cone constraints (see, e.g., Nesterov and Nemirovskii 1994).

## 6. Joint Pricing and Scheduling

We consider an extension of the model in §3 to maximize the expected profit by jointly optimizing the pricing of charging service fees and scheduling of EVs. We specify the formulation of this extended model in §6.1. We solve this model by leveraging the ECP approximations from §4 and consider the uncapacitated and capacitated cases in §6.2 and §6.3, respectively.

## 6.1. Problem Formulation

Service fee structure. We use one of the most common pricing schemes in practice, where the service fee is proportional to the quantity of electricity charged (also known as kWh-based) (Tesla 2020, Blink 2020). This means that in our model customers of type *v* pay *pu<sup>v</sup>* for the EV charging service, where  $p \in [p, \overline{p}]$  is the decision of unit service fee.

Price-dependent arrival rate. For ease of exposition and computational tractability, we assume that the customer arrival rate is linearly decreasing in the price:

$$
\lambda_v = \bar{\lambda}_v (1 - r_v p), \quad \forall v \in [V], \tag{16}
$$

where  $\bar{\lambda}_v$  is the maximum possible arrival rate of type *v* customers (e.g., when the EV charging service is free), and  $r_v > 0$  is the price sensitivity of customer type *v*. Consistent with the literature in the pricing of service operations (e.g., Mendelson and Whang 1990), this price-dependent arrival rate represents aggregating the individual customer's decision of whether to use the charging service according to their service valuations, which follow a uniform distribution  $F(\cdot)$  between  $[0, 1/r<sub>n</sub>]$ : Since only customers with valuation higher than *p* join the service at the charging station, the effective arrival rate of type *v* customers is given by  $\lambda_v = \bar{\lambda}_v(1 - F(p)) = \bar{\lambda}_v(1 - r_v p)$ .

Objective function. The objective is to jointly optimize the pricing decision *p* and the scheduling decision x to maximize the expected total profit, i.e., the total revenue minus total cost, as follows:

$$
\max_{p \in [\underline{p}, \overline{p}], \mathbf{x} \in \mathcal{X}} \quad \mathbb{E}_{\mathbb{P}^C} \left[ \sum_{v \in [V]} p u_v \tilde{z}_v - c(\mathbf{x}, \tilde{\mathbf{z}}) \right]. \tag{17}
$$

Note the problem (17) is much more challenging than the scheduling problem (3) because the underlying distribution  $\mathbb{P}^C$  in (17) depends on the pricing decision p. To solve (17) efficiently, we leverage our ECP approximations in §4 for both the uncapacitated and capacitated cases.

#### 6.2. The Uncapacitated Case

PROPOSITION 4. When  $C = \infty$ , the optimal value of JPS-U is a lower bound of (17):

$$
\sup_{\mathbf{x}\in\mathcal{X},p\in[\underline{p},\overline{p}],\mu>0,\lambda,\kappa,\gamma}\sum_{v\in[V]}pu_v\lambda_v-d(\kappa+\gamma)-\sum_{s\in[T]}e_sf_s(\mathbf{x},\lambda)
$$
\n
$$
\text{s.t.}\qquad(5a), (8), (16).
$$
\n(JPS-U)

The ECP formulation in Proposition 4 is labeled JPS-U, representing "joint pricing and scheduling" for the "uncapacitated" case. Note that unlike ECP-U, which is convex, JPS-U is non-convex as (8) defining JPS-U is non-convex due to the fact that p (and equivalently  $\lambda_v$ ) is a decision variable. Therefore, we propose an optimization procedure to solve JSP-U efficiently by alternating between fixing  $p$  and fixing  $x$  and  $\mu$ : Both of the resulting formulations are convex and can be easily solved using off-the-shelf solvers. Specifically, first, given p, the arrival rates  $\lambda_v$  are given in (16) and thus fixed, so we can follow Proposition 1 to reformulate JPS-U as the following convex optimization problem with only scheduling decisions:

$$
\sup_{\mathbf{x}\in\mathcal{X},\mu>0,\kappa,\gamma,\xi,\zeta}\sum_{v\in[V]}pu_v\lambda_v-d(\kappa+\gamma)-\sum_{s\in[T]}e_sf_s(\mathbf{x},\boldsymbol{\lambda})
$$
\ns.t. (5a) - (5d). (18)

Second, given  $x$  and  $\mu$ , JPS-U can also be reformulated as the following convex optimization problem as in Proposition 1:

$$
\sup_{p \in [\underline{p}, \overline{p}], \lambda, \kappa, \gamma, \zeta} \sum_{v \in [V]} u_v \overline{\lambda}_v (p - r_v p^2) - d(\kappa + \gamma) - \sum_{s \in [T]} e_s f_s(\boldsymbol{x}, \boldsymbol{\lambda})
$$
\n
$$
\text{s.t.} \quad (5a), (5d), (9), (16). \tag{19}
$$

Therefore, we can solve JSP-U by alternatingly solving problems (18) and (19) with the following algorithm, the convergence of which is shown in Proposition 5.

Algorithm 1: Alternating optimization for JPS-U

- 1. **Initialization** Set initial price  $p^{(0)} \in [p, \overline{p}]$ , iteration counter  $i \leftarrow 1$
- 2. Scheduling optimization Solve model (18) with input  $p^{(i-1)}$  and let  $x^*$  and  $\mu^*$  be the optimal solution; set  $x^{(i-1)} \leftarrow x^*$  and  $\mu^{(i-1)} \leftarrow \mu^*$ . Store optimal value  $val_1^{(i)}$ ;
- 3. Pricing optimization Solve model (19) with inputs  $x^{(i-1)}$  and  $\mu^{(i-1)}$ , and let  $p^*$  be the optimal solution; set  $p^{(i)} \leftarrow p^*$ . Store optimal value  $val_2^{(i)}$ ;
- 4. **Termination** If  $|val_2^{(i)} val_1^{(i)}| < \delta$  (where  $\delta > 0$  is a given small tolerance), set  $p^* \leftarrow p^{(i)}$ , solve model (18) with input  $p^{(i)}$  to obtain  $x^*$ , output  $p^*$  and  $x^*$ , and then stop. Otherwise, set  $i \leftarrow i+1$  and go back to Step 2.

**Output:** Pricing decision  $p^*$  and scheduling decision  $x^*$ 

PROPOSITION 5. *The sequence of optimal values*  $\{val_1^{(i)}, val_2^{(i)}\}$  *in Algorithm 1 is non-decreasing and converges to a finite value.*

#### 6.3. The Capacitated Case

Next, we develop an ECP approximation of model (17) for the capacitated case. We first derive a convex upper bound for the expected revenue in Proposition 6. Then this upper bound on the expected revenue minus the upper bound on the expected cost in Proposition 3 gives an approximation of the expected profit in (17). Note that in the capacitated case, there is only an approximation, different from the uncapacitated case where we obtain a lower bound.

Proposition 6. *The optimal value of the following convex optimization problem is an upper*  $bound\ of\ the\ expected\ revenue\ \mathbb{E}_{\mathbb{P}^C}\left[\sum_{v\in[V]}pu_v\tilde{z}_v\right]$ .

$$
\sup_{g\geq 0} \sum_{v\in [V]} g_v u_v
$$
  
s.t.  $\bar{\lambda}_v (p - r_v p^2) \geq g_v$   $\forall v \in [V]$  (20a)

$$
Cp \ge \sum_{v \in \mathcal{V}_t} g_v \qquad \qquad \forall t \in [T] \tag{20b}
$$

*Proof.* Suppose  $\mathbb{E}_{\mathbb{P}^C}[\tilde{z}] = \underline{\lambda}$ , then  $\mathbb{E}_{\mathbb{P}^C}\left[\sum_{v \in [V]} pu_v \tilde{z}_v\right] = \sum_{v \in [V]} pu_v \underline{\lambda}_v$ . We know  $\underline{\lambda} \leq \lambda$  due to arrival truncation in capacitated case. And  $\mathbb{P}^C[\tilde{z} \in \mathcal{Z}] = 1$  implies  $\sum_{v \in \mathcal{V}_t} \underline{\lambda}_v \leq C, \forall t \in [T]$ , from (1). Hence,

$$
\sup_{\boldsymbol{b}\geq\boldsymbol{0}}\left\{\sum_{v\in[V]}pu_vb_v:\boldsymbol{b}\leq\boldsymbol{\lambda},\sum_{v\in\mathcal{V}_t}b_v\leq C,\forall t\in[T]\right\}
$$
\n(21)

is an upper bound of  $\sum_{v \in [V]} pu_v \underline{\lambda}_v$  as  $\underline{\lambda}$  is feasible in problem (21). Next, we perform a change of variable by  $g_v = pb_v$ . Since  $p > 0$ , problem (21) is equivalent to

$$
\sup_{g\geq 0} \left\{ \sum_{v\in[V]} u_v g_v : g \leq \lambda p, \sum_{v\in\mathcal{V}_t} g_v \leq Cp, \forall t \in [T] \right\}.
$$
\n(22)

Substituting (16) into the model (22) concludes the proof.  $\Box$ 

Combining Propositions 6 and 3, which give tractable approximations of the expected revenue and cost, respectively, we obtain the following approximation of the expected profit:

$$
\sup_{\substack{\boldsymbol{x}\in\mathcal{X},\boldsymbol{p}\in[\underline{p},\overline{p}],\lambda,a,b\ge0,\boldsymbol{\nu}\ge0,y\ge0,\\ \boldsymbol{\mu}>0,\kappa,\gamma,\alpha,\beta\ge0,\boldsymbol{\xi},\boldsymbol{\zeta},\boldsymbol{\rho}\ge0\boldsymbol{g}\ge0}} \sum_{v\in[V]} g_v u_v - d(\kappa+\gamma+\alpha+\beta'\boldsymbol{\lambda}) - a - b'\boldsymbol{\lambda}
$$
\n
$$
\text{(JPS-C)}\tag{JPS-C}
$$
\n
$$
\text{s.t.}\qquad (13a) - (13h), (16), (20a) - (20b)
$$

As in the uncapacitated case, we can perform alternating optimization to efficiently solve problem JPS-C using an algorithm analogous to Algorithm 1: Fixing  $p$  (or equivalently  $\lambda$ ) in JPS-C, we obtain the scheduling optimization, which is an ECP; fixing all variables except  $p, g, \lambda, \gamma, \kappa$  in JPS-C, we obtain the pricing optimization, which is a conic program involving exponential cones and second-order cones. This algorithm has a similar convergence property as Algorithm 1 and is omitted for brevity.

## 7. Numerical Study

In this section, we examine the performance of our ECP approximation in benchmark with the SAA and DRO approaches using numerical instances calibrated to real data for both the main model and the extended model with the pricing decision. We also discuss how they perform considering several practical implementation issues. Lastly, we use ECP to generate managerial insights.

#### 7.1. Experiment Settings

We consider the planning horizon of a service provider to be one day as there are very few EVs charging across two consecutive days in public charging stations, as shown in our data below. Each period represents 15 minutes, which is used in practice to compute demand charges for most utility companies, such as Southern California Edison. Therefore, the number of periods *T* equals 96.

*Electricity tariff calibration.* We use Southern California Edison's electricity tariff *Schedule GS-2 General Service* (Neubauer and Simpson 2015), which is the tariff an electricity user such as an EV charging service provider is subject to. Under this tariff, we obtain the unit energy charge and demand charge depending on the hour of the day as follows:

$$
\hat{e}_t = \begin{cases}\n\$0.1466/kWh & \text{if } 13 \leq \lceil t/4 \rceil \leq 18 \text{ (on-peak hours)} \\
\$0.0895/kWh & \text{if } 9 \leq \lceil t/4 \rceil \leq 12 \text{ or } 19 \leq \lceil t/4 \rceil \leq 23 \text{ (mid-peak hours)} \\
\$0.0582/kWh & \text{otherwise (off-peak hours)}, \\
\hat{d}_t = \begin{cases}\n\$0.465/kW & \forall t \in [T] \neq [96] \text{ (all-period)} \\
\$0.540/kW & \text{if } 13 \leq \lceil t/4 \rceil \leq 18 \text{ (on-peak hours)} \\
\$0.165/kW & \text{if } 9 \leq \lceil t/4 \rceil \leq 12 \text{ or } 19 \leq \lceil t/4 \rceil \leq 23 \text{ (mid-peak hours)},\n\end{cases}
$$

where  $\lceil t/4 \rceil$  represents the hour index of period t (as one hour has four periods). Note that these demand charge parameters for the horizon of a day are obtained by dividing those for a month by

30 to maintain the same proportion of the demand charge relative to the total cost. More details on the calibration can be found in Online Appendix B. Note that the demand charge is the sum of allperiod demand charge (applicable to all periods), on-peak demand charge, and mid-peak demand charge: Though the models in §4 and §6 have only all-period demand charge for the expositional simplicity, the ECP models for the case with multiple demand charges are constructed analogously to those in §4 and §6.

*EV arrival rate calibration.* In this study, we use the public EV charging data from the U.K. Department for Transportation. While we acknowledge the discrepancies in EV charging behaviors across different markets, we use the U.K. charging events to approximate general EV arrival patterns in the computational experiments. The dataset in the U.K. contains all of the charging events from 27 local authorities in England from January 1 to December 31, 2017 (U.K. Department for Transport 2018). For each charging event, it provides the start time, end time, and electricity supplied. In particular, we consider the data of both rapid and fast chargers where the charging speed are above 22 kW and below 22 kW, respectively. The average charging duration for rapid chargers is short, about 35 minutes, and for fast chargers is long, about 4.7 hours. Though the value of using our model is higher for long-duration charging than for short-duration as the former gives more room for optimization, we focus on the latter instead due to the following reason: While ECP scales well and can be solved for both the short- and long-duration charging, SAA does not scale well and cannot be solved to a satisfactory level for long-duration cases. Hence, unless specified explicitly otherwise, all the results mentioned below pertain to the short-duration charging case. We demonstrate that the value of our model remains high even for short-duration charging. Limited experiments show that this value is much higher in the case of long-duration charging. We observe that the average charging speeds (electricity supplied divided by the stay duration, the time between arrival and departure) for most charging events are *not* at the maximum. This observation holds at the aggregate level across all locations, at the individual charging location (local authority) level, and at the individual charger level. This observation corroborates the need to optimize the scheduling of EV charging with various customer types, as studied in this paper. In the rest of the paper, we use the arrival pattern aggregated across all locations unless specified otherwise.

After pre-processing the dataset (see details in Online Appendix B), we obtain an average of 278.65 daily charging events. We then discretize time into 15-minute (shortened as 15-min) intervals. In particular, we count any partial arrival/departure period as a full period: For example, a car that arrives at 12:56 and leaves at 13:22 is treated as staying between 12:45 to 13:30, i.e., it







(a) Normalized histogram of stay duration (b) Normalized histogram of electricity charged

stays for three 15-min periods<sup>5</sup>. Then we compute all policies based on this 15-min discretization. We can easily convert this 15-min policy to a policy that is implementable on a 1-min discretization using a simple translation algorithm (see more details in §7.4.1). We show that the performance of this converted policy is close to that of the policy computed on the 1-min discretization, and the performance comparison between the converted ECP and converted SAA is the same as that on 15 min discretization. In addition, there are two more reasons to use 15-min discretization: First, it is consistent with the measurement for the demand charge used by utility firms in the U.S., and thus often used in the literature (Glassmire et al. 2012, Zhang et al. 2017, and Zhang and Augenbroe 2018); second, finer time discretization levels are computationally more challenging due to memory and CPU time limitations, for instance, the number of random variables and decision variables under 1-min discretization is  $15<sup>2</sup>$  and  $15<sup>3</sup>$ , respectively, times of those under 15-min discretization.

We plot in Figure 1 the resulting density (or normalized frequency) of EV stay duration (time between arrival and departure) and electricity charged. The mean (standard deviation) of stay duration is 35.83 (27.92) minutes; the mean (standard deviation) of electricity charged is 10.69 (7.64) kWh. We then estimate the arrival rate of each type as the average number of charging events *of each type* per day within the entire horizon. Note that as our data spans a year, *each* type has 365 data points for estimating its arrival rate: For each day, each type has *one realization* of its arrivals, even if the number of arrivals on that day is zero. Subsequently, among about 80*,* 000 customer types with distinct start times, departure times, and total electricity charged, we identify 14*,* 284 customer types with positive arrival rates. We use these arrival rates for the *given* truncated

<sup>5</sup> We also tried time discretization by rounding up the stay duration to the integer multiple of 15 minutes, i.e., [stay duration/15]  $\times$  15. We find that it performs worse as it gives a higher total expected cost for both ECP and SAA.

Poisson distributions in our SP model to compare the performances of all policies. While we assume Poisson arrivals with known rates, there may be estimation errors in arrival rates due to insufficient data and model mis-specification of the arrival distribution in practice, and we therefore consider these two cases in §7.4.2.

Next, we consider managing a hypothetical charging station to charge EVs. We plot in Figure 2 the expected number of EVs at the charging station in a 1-day horizon under various values of the number of chargers  $C$  to achieve different service levels (the probability that a customer finds a charger upon arrival). As seen, the expected number of customers at the station follows a daily pattern: It starts low from midnight to 6 am, increases significantly to the peak at around 1 pm, and finally decreases significantly until midnight. When *C* = 30, there is almost no lost customer (with service level 99.86%): The corresponding line in Figure 2 almost coincides with that for the uncapacitated case. However, with  $C = 25, 20,$  and 15, the service level drops quickly, to  $98.83\%$ , 94.07%, and 82.54%, respectively.

#### Figure 2 (Color Online) Average number of EVs at the charging station



*Other parameters.* We use the specifications of Nissan Leaf, a popular customer type. Specifically, we set  $\hat{U} = 62$  kWh using the energy capacity of Nissan Leaf (Nissan 2019);  $\hat{K} = 10.75$  kWh/period based on the rapid AC chargers with power at 43 kW (Zap-Map 2019); and  $\hat{\eta} = 0.9$  using the efficiency of lithium-ion battery in Nissan Leaf (Karlsson and Kushnir 2013).

Note that our 96-period instance is large-scale: There are in total 80,736 random variables and 692*,* 702 decision variables. In this case, the benchmark DRO-C cannot be solved, and thus we focus on comparing ECP with SAA in the rest of this section. We also consider a numerical setup with small problem instances where there are only 6 periods and DRO-C can be solved. See Online Appendix C for more details of the comparisons among all three approaches.

#### 7.2. Computation

We implement SAA using the following two methods: Solving the extensive form by the interior point (also known as barrier) method as well as the L-shaped method (also known as Benders decomposition) with either single cut or multi-cut (Birge and Louveaux 2011). In the L-shaped method, we first generate optimality cuts (as our problem has complete recourse) by solving *N* subproblems over a simplex and then add these cuts to the master problem. In our multi-cut implementation, we do not add generated cuts for all the scenarios in each iteration, as the master problem (where the computational burden lies) is so large that we can only perform four iterations within 12 hours (our specified maximum computation time). Instead, we add into the master problem in each iteration the 100 cuts which have the largest gaps between the recourse function and its piecewise affine lower bound at the current solution. We use dual simplex without presolve to solve all the linear programming problems in the L-shaped method.

As in our SP model the arrival distribution is *given*, we compute and evaluate ECP and SAA using the same calibrated arrival distribution. In particular, we generate 8,000 sample paths to compute the actions  $x^S$  under SAA. (Note that *computing* the action  $x^E$  under ECP does not require samples; only *evaluating* the action under ECP does.) After solving for  $x<sup>E</sup>$  and  $x<sup>S</sup>$  using the calibrated arrival distributions and other parameters, we generate from the same distribution via Monte Carlo simulation 10,000 samples to evaluate both, sufficient to show statistically significant performance comparison between both approaches. (The results do not change when using more samples in the evaluation step.) We evaluate  $x^E$  and  $x^S$  on these samples by computing the mean and standard deviation of  $c(x^E, \tilde{z})$  and  $c(x^S, \tilde{z})$ , respectively. We implement SAA five times as the solutions obtained from SAA are sample-dependent and report the final performance as the average of these five instances.

All of our numerical experiments are conducted in Python while using the MOSEK 9.2 solver on a Windows-OS 64-bit with 24 GB of RAM and an Intel i7-8750H CPU@2.20GHz Processor. All the problem instances except the L-shaped method are solved via the interior point method. For all methods, we set both primal and dual feasibility tolerance  $10^{-8}$  and relative duality gap  $10^{-8}$ . When the solver cannot find solutions to this accuracy, we accept the solution if it is feasible but not necessarily optimal.

#### 7.3. Comparing ECP with SAA

Table 1 compares under different values of  $C$  the performances of ECP and SAA. In this table, SAA uses the interior point method, as it is faster than the L-shape method with either single cut or multi-cut (see such experiments on small, six-period instances in Online Appendix C). As mentioned in §7.2, SAA uses 8*,* 000 samples in computation, as SAA with a larger sample size cannot be solved within 12 hours (see more details on SAA with different sample sizes in Figure 11 in Online Appendix D). This table shows that compared with SAA, ECP results in not only a lower standard deviation of the mean cost but also a lower mean cost itself. (Note that this difference in mean cost is significant compared to the standard error, which is the standard deviation of the total cost divided by the square root of the evaluation sample size of 10,000.) For instance, when *C* equals 30, ECP-C has a standard deviation and mean cost of 67.55 and 805, respectively, lower than those of SAA, which are 69.94 and 816.39, respectively. That is, the robustness of ECP-C (or lower standard deviation) does not come at the cost of a higher mean cost, contrary to known tradeoff (e.g., Choi and Ruszczyński 2008, Hao et al. 2020). Similar results are also observed when comparing ECP with SAA using each arrival pattern from the three charging locations (local authorities) that have the largest number of charging events in the U.K. data.<sup>6</sup> ECP-C beats SAA on robustness as expected because ECP-C leverages the ambiguity set inspired by DRO approaches (see §4.2). Compared to ECP models, SAA results in a high mean cost because it does not scale well and is thus hard to solve to optimality on large-scale instances.

C	Method			CPU	
		Total cost			
		Mean	Std	Time(s)	
15	ECP-C	600.40	35.20	312.33	
	SAA	608.91	36.76	13015.66	
20	ECP-C	717.73	44.42	354.66	
	SAA	727.38	46.22	15806.33	
25	ECP-C	783.66	57.13	445.66	
	SAA	794.04	59.44	16207.17	
30	ECP-C	805.00	67.55	320.56	
	SAA	816.39	69.94	16005.50	
$\infty$	ECP-U	808.90	71.30	194.22	
	SA A	819.80	73.64	17673.80	

Table 1 Performance comparison between ECP-C (or ECP-U) and SAA

To illustrate the reason why SAA has a higher mean cost than ECP, Figure 3 plots the extrapolation on the objective value of SAA (i.e., that of (14)) and the evaluation of SAA (i.e.,  $\pi^S$ ) at different sample sizes given  $C = 30$ : The number of samples required for  $\pi^S$  to be lower than  $\pi^E$ at  $C = 30$  (which is 805, see Table 1) is large (about 23,000); that required for solving SAA to

<sup>6</sup> Details are omitted for brevity and are available upon request.



Figure 3 (Color Online) Expected cost under SAA at different sample sizes given  $C = 30$ 

Figure 4 (Color Online) Objective value and total CPU time of solving SAA with 38,000 samples using the **L-shaped method with multi-cut given**  $C = 30$ 



optimality is even larger (about 38,000), resulting in an estimated optimal mean cost of 799.30. However, this mean cost is not attainable because solving SAA to near-optimality on our instances with 38,000 samples is computation-prohibitive, where we use the L-shaped method with multi-cut (as the interior method cannot be used due to memory issues). To demonstrate that, Figure 4 shows that SAA cannot converge even within 36 hours: Figure 4(a) plots both the lower bound of the objective value (obtained as the optimal value of the master problem) and the upper bound (obtained as the objective value of SAA at current solution), and Figure 4(b) shows that the CPU time per iteration increases exponentially when the number of iteration is large. The SAA solution using 38,000 samples at 36 hours leads to a mean cost of 866.06 and a standard deviation of 77.21, larger than those of SAA in Table 1.

Table 1 also lists the CPU computation time (excluding the evaluation time) of ECP and SAA in the last column. As seen, ECP computes about 50 times faster than SAA, suggesting that ECP is a more efficient computational approach to solve large-scale instances.

The superior performance of ECP over SAA is even more pronounced on numerical instances of long-duration charging, about 4.7 hours with fast chargers (see §7.1). Table 2 displays the mean and standard deviation of the total cost under both approaches, which shows that both measures under ECP-C are about 10% lower than those of SAA. This is mainly because SAA does not scale well, so we can compute SAA with only 1,000 sample paths due to memory limit: The number of customer types is much larger in the long-duration charging case, so we can no longer use the same number of sample paths (i.e., 8,000) for computing  $x<sup>S</sup>$  from SAA as in the short-duration charging case. (Note that when evaluating both ECP-C and SAA, we continue using 10,000 sample paths from the arrival distributions calibrated for long-duration instances.) As such, we use short-duration charging cases in the rest of the paper.

C	Method	Total cost		$\mathbb{E}\{$ Demand charge
		Mean Std		$/$ total cost}
60	$ECP-C$	404.33	31.52	0.58
	SAA(1,000)	454.71	38.31	0.64
50 40	$ECP-C$	372.67	27.93	0.58
	SAA(1,000)	417.55	34.53	0.64
	ECP-C	327.68	25.78	0.58
	SAA(1,000)	350.92	29.93	0.62

Table 2 Performances of ECP-C and SAA on long-duration instances

*Notes.* The number in the parenthesis after SAA indicates the number of sample paths used to compute SAA.

We next show that ECP is near optimal as its optimality gap is at most around  $4\%$  in all instances. We estimate the optimality gap of ECP (i.e.,  $(\pi^E - \pi^*)/\pi^* \times 100\%)$  using the optimal value of SAA (not listed in Table 1) as a lower bound of  $\pi^*$  (Shapiro et al. 2009). For instance, when  $C = 30$ , this optimality gap of ECP-C is at most  $(805.00 - 773.67)/773.67 \times 100\% \approx 4.05\%$ , where  $\pi^{E}$  is 805.00 (see Table 1) and the optimal value of SAA given sample size 8,000 is 773.67. If we use the estimated optimal value 799.30 in Figure 3, the gap decreases to only 0.71%. Note that we can also similarly estimate the optimality gap of ECP-U to be at most 4.24%. This gap is much smaller than that computed using the theoretical bound in (12), 214.67, divided by  $\pi^E = 808.90$ (see Table 1), which gives the optimality gap of 26*.*5%, as the former optimality gap is computed *after* solving ECP-U while the latter is computed *before* solving it.

In sum, ECP performs better than other common approaches and achieves a near-optimal total cost. In  $\S 7.4$ , we consider different practical implementation issues in the numerical study. In  $\S 7.5$ , we also use ECP to generate managerial insights to demonstrate its practical relevance further.

#### 7.4. Practical Considerations in Implementation

7.4.1. Time discretization. As mentioned in §7.1, we discretize time into increments of 15 minutes and compute all policies on this discretization, and cannot discretize time into finer discretization, such as 1-min, due to multiple reasons including CPU and memory limitations. We next examine the effect of our discretization by comparing the performance of our 15-min policies with those on 1-min discretization. To this end, we have to first convert our 15-min policies to be implementable (i.e., feasible) on 1-min discretization. Note that if we simply divide the charging quantity for each type in each 15-min period by the number of 1-min's within this period, the resulting 1-min level charging quantities in the arrival and departure periods are not necessarily feasible as they may violate the corresponding charging speed limit constraint in  $\mathcal{X}$ . For example, if the 15-min policy in the arrival period is to charge at the maximum speed (i.e., the charging quantity is  $K/\eta$ , the 1-min policy is to charge  $K/(S\eta)$ , where *S* is the number of 1-min intervals in the arrival period: If  $S < 15$ , this exceeds  $K/(15\eta)$ , the charging speed limit under 1-min. Thus, before the division step, we have to shift the charging quantity on both the arrival and departure periods to the middle periods of the charging time window.

Specifically, we develop Algorithm 2. Note that in this algorithm, if  $s_v = \tau_v$  (i.e., the arrival period is the same as the departure period), the division step (on Line 22) is sufficient to convert the  $15$ min policy to be 1-min implementable. Hence, the algorithm needs to shift only when  $s_v < \tau_v$  (see Line 7 onward). Lines 8-13 (14-19) shift the excess charging quantity from the arrival (departure) period to the middle periods of the charging time window. To demonstrate the effectiveness of Algorithm 2, we compare the following two policies:

- Opt-15minTo1min policy: the 1-min charging policy transformed via Algorithm 2 from the *optimal* charging policy solved on 15-min discretization.
- Opt-1-min policy: the *optimal* charging policy solved directly on 1-min discretization.

Using the optimal policies allows us to isolate the effect of approximation and focus on the performance of Algorithm 2 itself. Therefore, we have to use small 1-min numerical instances where we can obtain optimal policies by solving SAA with sufficiently many samples. In particular, we choose a one-hour horizon (i.e., 5 am-6 am), where the number of types with positive arrival rates is 2, 439. On such instances, 8,000 samples are sufficient for SAA to generate optimal policies for the SP. Our experiments on other one-hour horizons with a similarly small number of types show the same qualitative results, which are omitted for brevity. For obtaining the Opt-15minTo1min policy, we first compute the optimal 15-min policy under SAA by generating 8,000 samples from Poisson distributions of customer arrivals with 15-min discretization and then convert it using Algorithm 2. For obtaining the Opt-1-min policy under SAA, we use 8,000 samples generated from Poisson Algorithm 2: Transform a 15-min charging policy *x* into a 1-min implementable policy *y* Input: 15-min charging policy *xv,t*

1 for 1-min customer type  $v' \in [V']$  do  $2 \mid \text{set } s_v \leftarrow [s_{v'}/15], \tau_v \leftarrow [\tau_{v'}/15], v \leftarrow (s_v, \tau_v, u_{v'});$ 3 for 15-min period  $t \in \{s_v, ..., \tau_v\}$  do 4 \ \ \ \ \ compute the actual 1-min stay duration  $S_t$  of EV of type  $v'$  in period t;  $\mathbf{5}$  | compute  $\bar{x}_{v,t} \leftarrow K/(15\eta) \cdot S_t;$  $6 \mid end$  $\tau$  | if  $s_v < \tau_v$  then 8 compute the excess charging quantity in *arrival* period  $E \leftarrow \max\{0, x_{v,s_v} - \bar{x}_{v,s_v}\};$ 9 | update  $x_{v,s_v} \leftarrow x_{v,s_v} - E$ ; set  $t \leftarrow s_v + 1$ ; 10 | while  $E > 0$  do 11 | update  $x_{v,t} \leftarrow x_{v,t} + E$ ; compute  $E \leftarrow \max\{0, x_{v,t} - \bar{x}_{v,t}\}$ ; 12 update  $x_{v,t} \leftarrow x_{v,t} - E$ ; set  $t \leftarrow t + 1$ ;<br>
13 end end 14 compute the excess charging quantity in *departure* period  $E \leftarrow \max\{0, x_{v,\tau_v} - \bar{x}_{v,\tau_v}\}$ ; 15 | update  $x_{v,\tau_v} \leftarrow x_{v,\tau_v} - E$ ; set  $t \leftarrow \tau_v - 1$ ; 16 | while  $E > 0$  do 17 | update  $x_{v,t} \leftarrow x_{v,t} + E$ ; compute  $E \leftarrow \max\{0, x_{v,t} - \bar{x}_{v,t}\}$ ; 18 | update  $x_{v,t} \leftarrow x_{v,t} - E$ ; set  $t \leftarrow t - 1$ ; 19 | end 20 end  $21$  for *1-min period*  $t' \in \{s_{v'},..., \tau_{v'}\}$  do 22  $|$  set  $t \leftarrow \lceil t'/15 \rceil$ ; set  $y_{v',t'} \leftarrow x_{v,t}/S_t$ ; 23 end 24 end

**Output:** 15minTo1min implementable charging policy  $y_{v',t'}$ 

distributions of customer arrivals with 1-min discretization. (Note that our SP model on 15-min discretization can be easily extended to obtain one on 1-min discretization.) We then evaluate both the Opt-15minTo1min and Opt-1-min policies on new 10,000 samples generated from the same Poisson distributions with 1-min discretization. We display the mean and standard deviation of the total cost from both policies in Table 3.

As seen from Table 3, the mean total costs between the Opt-15minTo1min and Opt-1-min policies are close. This shows that our Algorithm 2 is effective in converting a 15-min policy to a 1-min implementable policy without losing much optimality.

Hour	Policy		Total cost	Ratio of
		Mean   Std		demand charge
$5\mathrm{am}\text{-}6\mathrm{am}$	$Opt-1-min$	1.046	1.144	0.361
	Opt-15minTo1min   $1.061$   $1.163$			0.370

Table 3 Performance over a one-hour time horizon (5 am-6 am) with 1-min discretization





*Notes.* The number in the parenthesis after SAA indicates the number of samples used to compute SAA.

Lastly, we show that if we use Algorithm 2 to convert ECP and SAA policies obtained on 15-min discretization to their respective 1-min policies, their performance comparison is the same as that under 15-min discretization. We focus on one-hour horizons where the customer types with positive arrival rates are large, which means the computation is intensive. In particular, we choose 8 am-9 am and 12 pm-1 pm, where the number of types is 16*,* 633 and 23*,* 086, respectively. Consequently, SAA cannot be computed with 8,000 samples due to memory limit, and we have to use 3,000 and 2,000 computation samples, respectively. In contrast, ECP can be solved easily as it scales well and does not require samples in computation. (Note that we continue to use 10,000 sample paths for evaluating both policies.) We display the mean and standard deviation of the total expected costs under both ECP-15minTo1min and SAA-15minTo1min in Table 4. As seen, similar to the result on 15-min discretization in Table 1, in either hour, ECP results in both a lower mean and a lower standard deviation of the total cost than SAA.

In sum, our Algorithm 2 is effective in converting any 15-min policy to a 1-min policy, and the performance comparison of these converted policies is consistent with those on 15-min discretization. Therefore, due to the computational challenges of 1-min discretization, we focus on 15-min discretization in our numerical experiments.

7.4.2. Data-driven settings. Though in our SP model we assume Poisson arrival distributions with known arrival rates, we next consider the data-driven setting where there may be estimation errors in the arrival rate (due to various reasons including insufficient data) or misspecification of arrival distribution (e.g., which may not be truncated Poisson). We consider in such

settings whether ECP or SAA is more affected by comparing their out-of-sample performances. We benchmark them with two other approaches. The first approach is empirical risk minimization (ERM), which minimizes the empirical cost on training samples. It differs from SAA used in our paper in that ERM directly uses the training samples to calculate the empirical cost, whereas SAA approximates the expected cost by generating new samples from the truncated Poisson distributions with calibrated arrival rates. The second approach is a robust version of the ECP approach, labeled RECP. Under this approach, each arrival rate is assumed to be in an interval centered at the estimated arrival rate, and then we minimize the worst-case total expected cost, which occurs when the arrival rate is the largest as the cost is non-decreasing in each arrival rate. Specifically, if we denote the ECP model as  $\min_{x \in \mathcal{X}} g^E(x, \lambda)$ , the RECP model is obtained by allowing  $\lambda$  to be in the interval  $[\underline{\lambda}, \overline{\lambda}]$ , where  $\underline{\lambda}$  and  $\overline{\lambda}$  are the given lower and upper bounds of  $\lambda$ , respectively. Then we have

$$
\min_{\boldsymbol{x}\in\mathcal{X}}\max_{\boldsymbol{\underline{\lambda}}\leq\boldsymbol{\lambda}\leq\overline{\boldsymbol{\lambda}}}g^{E}(\boldsymbol{x},\boldsymbol{\lambda})\Longleftrightarrow\min_{\boldsymbol{x}\in\mathcal{X}}g^{E}(\boldsymbol{x},\overline{\boldsymbol{\lambda}}),
$$

where the equivalence of the two formulations is due to the fact that  $g^{E}(\boldsymbol{x}, \boldsymbol{\lambda})$  is non-decreasing in  $\lambda$ . We choose  $\overline{\lambda} = \hat{\lambda} + M \cdot \left( \frac{\sum_{v \in [V]} \hat{\lambda}_v}{\sqrt{N}V} \right) \mathbf{1}$ , where  $\hat{\lambda}$  is the estimated arrival rates, *M* is a hyperparameter, which we set as 0*.*5 based on cross-validation, and *N* is the number of training samples. Estimation errors in arrival rates. We first examine how estimation errors in arrival rates affect the performances of all policies where the arrival distribution is correct. To this end, we use the same distribution (i.e., truncated Poisson) for generating training and testing samples. In particular, we generate  $N \in \{50, 100, 200, 365\}$  i.i.d. training samples (recall 365 is the number of data points for calibrating the arrival rate of each type in our numerical study), which we use to estimate arrival rates  $\lambda$ . As N increases, there is more data and thus a smaller estimation error. We then feed these estimated arrival rates into ECP, SAA, and RECP to compute their respective policy actions. Note that SAA uses  $\lambda$  to generate 8,000 new samples to compute its actions, while ERM directly uses the *N* training samples as inputs for the optimization. We then generate from the original truncated Poisson distribution 10,000 testing samples to compute the out-of-sample performances of all policies.

The results are shown in Table 5 for  $C = 30$ . (Results for when  $C = 15$ , 20, and 25 are qualitatively similar, which are omitted for brevity.) Among all policies, ERM has the highest total expected cost and the highest associated standard deviation and thus performs the worst. SAA improves significantly on ERM, as different from ERM it can generate new samples to represent scenarios other than the training samples. ECP performs better than both ERM and SAA for any training sample size. This shows that our ECP continues to perform better than SAA (and a related approach ERM) in the presence of estimation errors. The performance of ECP can be further

N	Method	Total cost		Ratio of
		Mean	Std	demand charge
	ECP	835.39	72.97	0.56
50	<b>ERM</b>	876.84	80.20	0.58
	SA A	839.45	73.83	0.56
	RECP	804.88	67.65	0.55
	ECP	823.28	71.12	0.56
100	<b>ERM</b>	875.77	79.49	0.58
	SA A	829.56	72.32	0.56
	<b>RECP</b>	804.14	67.47	0.55
	<b>ECP</b>	815.55	69.86	0.55
200	<b>ERM</b>	867.05	77.93	0.58
	SA A	823.58	71.14	0.56
	RECP	803.81	67.36	0.55
365	<b>ECP</b>	810.29	68.78	0.55
	<b>ERM</b>	857.30	76.10	0.57
	SAA	820.46	70.59	0.56
	RECP	803.58	67.27	$0.55\,$

Table 5 Policy performances in the presence of estimation errors when *C* = 30

improved if we add some robustness to it, i.e., using RECP, which significantly improves on ECP, especially when the training sample size  $N$  is small (for instance, when  $N = 50$  or 100).

To assess how the presence of estimation errors affects the performance of each policy compared to the case without estimation error, we compare Table 5 with the row for  $C = 30$  in Table 1. We see that the presence of estimation errors undermines the performances of both ECP and SAA as the total expected costs for each policy given any *N* in Table 5 are higher than those in the row for *C* = 30 in Table 1 as expected, especially for small *N*. However, when *N* = 365 (again the number of data points for calibrating the arrival rate of each type in our numerical study), the mean total costs for ECP and SAA are 810.29 and 820.46, respectively, which are close to those of ECP and SAA with accurate arrival rates in Table 1, i.e., 805 and 816.39, respectively. This indicates that though our SP model (where ECP and SAA derive their formulation from) has a large number of customer types and thus may seem over-parameterized, the performance of this model is satisfactory even in the presence of estimation errors.

Model mis-specification of arrival distribution. We next consider the case when there is model mis-specification, i.e., the arrival distribution may not be truncated Poisson. To this end, we use our real arrival data, where the distribution is not necessarily Poisson, and evaluate the performances of all policies to see which one is affected the most.

We repeat the following experiment five times. In each experiment, we randomly equally partition the 365 data points (each corresponding to a day) into 5 subsamples, where 4 subsamples are used

No. of	Method	Total cost		Ratio of
experiment		Mean	$\operatorname{Std}$	demand charge
	ECP	809.10	102.53	$0.55\,$
$\mathbf{1}$	ERM	851.41	108.06	0.57
	SAA	818.18	103.24	0.56
	<b>RECP</b>	792.59	95.92	$0.54\,$
	ECP	808.65	101.64	0.55
$\overline{2}$	ERM	850.69	107.33	0.57
	SAA	818.14	102.36	0.56
	<b>RECP</b>	792.31	94.55	0.54
	ECP	808.89	101.38	0.55
3	ERM	854.26	108.25	0.58
	SAA	815.82	101.21	$0.55\,$
	<b>RECP</b>	792.44	95.74	0.54
	ECP	808.55	101.07	0.55
$\overline{4}$	ERM	851.60	110.38	0.57
	SAA	816.48	103.44	$0.55\,$
	<b>RECP</b>	792.53	95.25	0.54
	ECP	809.37	102.32	$0.55\,$
$\bf 5$	ERM	853.07	107.42	$0.57\,$
	SAA	816.54	103.11	$0.56\,$
	RECP	792.55	96.18	0.54

Table 6 Policy performances with model mis-specification when randomly partitioning real data at *C* = 30

for training and the remaining one for testing. We use the training subsamples to estimate the arrival rates and feed the estimates into ECP, SAA, and RECP approaches. (Again, ERM directly uses the training subsamples without estimating the arrival rates.) Then we evaluate the solutions obtained for each policy on the testing subsample and report their out-of-sample performances.

We present the results for each experiment given  $C = 30$  in Table 6 (those for other *C* values are omitted for brevity). As seen, the results are qualitatively very similar to the case with estimation errors in arrival rate. In particular, ERM performs the worst, while SAA improves significantly on ERM. ECP performs better than both ERM and SAA. RECP performs the best. These qualitative results continue to hold if we partition real data into training and testing in chronological order for given different numbers of chargers (see results in Table 7), i.e., when we use the first  $4/5$  of the entire data set for training and the rest for testing.

In sum, we show that in the presence of arrival rate estimation errors or arrival distribution mis-specification, our ECP approach continues to outperform SAA and a related approach (ERM). The performance of ECP can be significantly improved if we add simple robustness to it. All these results demonstrate the practical use of ECP in data-driven settings.

$\overline{C}$	Method	Total cost		Ratio of
		Mean	$\operatorname{Std}$	demand charge
	ECP	794.89	76.14	0.56
30	<b>ERM</b>	831.82	85.90	0.58
	SAA	805.14	72.85	0.57
	RECP	774.37	64.35	0.55
	ECP	776.61	71.46	0.56
25	<b>ERM</b>	819.65	82.42	0.58
	SAA	786.99	72.97	0.56
	<b>RECP</b>	757.07	58.54	0.55
	ECP	718.94	58.13	0.54
20	<b>ERM</b>	763.17	61.79	0.57
	SAA	729.55	54.97	0.55
	<b>RECP</b>	699.37	50.36	0.53
15	ECP	609.64	42.91	0.53
	<b>ERM</b>	657.45	43.06	0.57
	SAA	616.17	42.07	0.54
	RECP	587.95	40.68	0.52

Table 7 Policy performances with model mis-specification when partitioning real data in chronological order

7.4.3. Adaptive EV charging setting. As our ECP policy is static and menu-based that is tailored to customer types, we next examine how well it performs in an adaptive setting where the uncertainty is revealed over time. We benchmark it with a practically-implementable adaptive equal charge policy and the rolling horizon implementation of an LP and our ECP.

Equal charge policy: This policy is motivated by Blink and ChargePoint. For instance, Blink Charging (2022) "allows equal output to each charger based on the number of stations being used at one time. When one EV is charging, the EV will receive the maximum output. When others connect, the load will be equally shared between them." ChargePoint (2022) uses a very similar policy. We implement their "equal charge" (EQC) policy to ensure the feasibility constraints in *X* by first determining the total charging rate  $X \geq 0$  for all periods and then setting in period t the charging quantity  $x_{v,t}$  for each  $v \in V_t$  as

$$
x_{v,t} = \min\left\{K, u_v - \eta \sum_{\ell=s_v}^{t-1} x_{v,\ell}, \max\left\{\frac{\eta X}{\sum_{v \in \mathcal{V}_t} z_v}, u_v - \eta \sum_{\ell=s_v}^{t-1} x_{v,\ell} - K(\tau_v - t)\right\}\right\}/\eta. \tag{23}
$$

(23) shows that  $\eta x_{v,t}$  is no larger than any of the three terms: The charging speed limit *K* (first term), the remaining charging requirement  $u_v - \eta \sum_{\ell=s_v}^{t-1} x_{v,\ell}$  (second term), and the third term. The third term itself is the max of two expressions: The first distributes equally the total charging rate  $\eta X$  to all the EVs available for charging at the station (i.e.,  $\sum_{v \in V_t} z_v$ ); the second,  $u_v$  $\eta \sum_{\ell=s_v}^{t-1} x_{v,\ell} - K(\tau_v - t)$ , is the minimum charging quantity in period t to ensure satisfying the charging requirement before the departure time: The difference between the remaining charging

requirement (i.e.,  $u_v - \eta \sum_{\ell=s_v}^{t-1} x_{v,\ell}$ ) and the maximum possible quantity that can be charged in the remaining periods (i.e.,  $K(\tau_v - t)$ , which is the maximum speed times the number of periods left). Note that when *X* is large enough, EQC becomes the "maximum-speed charging" policy, which we discuss in §7.5.1. Lemma 2 below shows that this EQC policy is feasible, i.e.,  $x \in \mathcal{X}$ . In the experiment, we adopt a golden-section search procedure to determine the optimal *X* over the interval  $[0, X]$ , where *X* is a large number.

### LEMMA 2. *Given*  $X \geq 0$ *, for*  $x$  *under the EQC policy specified by* (23)*, we have*  $x \in \mathcal{X}$ *.*

Rolling horizon implementation of LP and ECP: We implement the rolling horizon version of a linear program (RH-LP) and our ECP (RH-ECP) similar to Lee et al. (2020). Specifically, in each period, given the system's status (e.g., the number of EVs at the station and remaining charging requirements), RH-LP resolves the deterministic counterpart of our SP model until the end of the horizon by replacing the future arrivals with their mean values. This deterministic optimization can be rewritten equivalently as an LP. Please find more details and our Algorithm 3 to implement RH-LP in Online Appendix E.1. We implement RH-ECP in a similar rolling fashion and relegate the details and our algorithm to Online Appendix E.2.

We next compare ECP with all three benchmark policies. In particular, we compare ECP with EQC on the large-scale (96-period) numerical instances, and with RH-LP and RH-ECP on the small-scale (six-period) instances (discussed in Online Appendix C). This is because even though both RH-LP and RH-ECP can be *computed* efficiently on large-scale instances and are thus practically implementable, it is computational-prohibitive to *evaluate* both RH-LP and RH-ECP on these large-scale instances. Both are practically implementable because the computation time in each period for RH-LP and RH-ECP, respectively, is 3.75 seconds and 2.8 minutes on average (with 5 minutes at maximum), both of which are sufficient for most practical purposes as they only need to resolve every 15 minutes or longer. But in terms of evaluation time, recall that we have to evaluate their performances over 10,000 sample paths, and the evaluation on *one* sample path already takes about 3.75 sec/period  $\times$  96 periods = 6 minutes and 2.8 min/period  $\times$  96 periods  $\approx$  4.5 hours, respectively: This is because both approaches need to resolve in each period with the realized uncertainty.

We first show the performances of ECP and EQC evaluated with 10,000 samples in Table 8. Although ECP is static and menu-based while EQC is adaptive based on the realized uncertainty, we find ECP outperforms EQC on each instance. This is because, while EQC optimizes only over the total load *X*, ECP optimizes the entire charging schedule tailored to each customer type. The

$\mathcal{C}$	Method	Total cost		$\mathbb{E}\left\{\right.$ Demand charge
		Mean	Std	$/total cost$ }
30	<b>ECP</b>	805.00	67.55	0.55
	EQC	810.54	69.93	0.55
25	<b>ECP</b>	783.66	57.13	0.54
	EQC	791.32	60.74	0.54
20	<b>ECP</b>	717.73	44.42	0.53
	$_{\mathrm{EQC}}$	731.87	49.09	0.53
15	<b>ECP</b>	600.40	35.20	0.51
	EQC	620.20	39.59	0.52
10	ECP	441.34	27.71	0.50
	EQC	462.77	31.19	0.52
5	ECP	250.43	19.43	0.51
	$_{\mathrm{EQC}}$	267.75	21.73	0.54

Table 8 Performances of ECP and EQC on large-scale (96-period) charging instances

Table 9 Performances of ECP, RH-LP, and RH-ECP on small-scale (6-period) instances

$\overline{C}$	$\boldsymbol{d}$	Method	Total cost		$\mathbb{E}\left\{\text{Demand charge}\right\}$
			Mean	Std	$/total cost$ }
		$\overline{ECP}$	6.00	1.76	0.43
	0.23	RH-LP	6.02	1.76	0.46
		RH-ECP	5.99	1.72	0.43
		$\overline{ECP}$	8.64	2.53	0.60
12	0.46	RH-LP	8.81	2.63	0.63
		RH-ECP	8.59	2.46	0.60
		ECP	13.91	4.08	0.75
	0.93	RH-LP	14.40	4.40	0.77
		RH-ECP	13.77	3.93	0.75
		<b>ECP</b>	5.75	1.52	0.44
	0.23	RH-LP	5.77	1.53	0.46
		RH-ECP	5.71	1.48	0.43
	0.46	ECP	8.28	2.18	0.60
10		RH-LP	8.44	2.29	0.63
		RH-ECP	8.18	2.10	0.60
		ECP	13.49	3.56	0.75
	0.93	$RH-LP$	13.80	3.85	0.77
		RH-ECP	13.12	3.37	0.75
		<b>ECP</b>	5.41	1.25	0.43
	0.23	RH-LP	5.26	1.22	0.47
		RH-ECP	5.23	1.19	0.43
		$\overline{ECP}$	8.08	1.84	0.61
8	0.46	RH-LP	7.72	1.83	0.64
		RH-ECP	7.59	1.69	0.60
		ECP	13.00	2.95	0.76
	0.93	RH-LP	12.65	3.09	0.78
		RH-ECP	12.15	2.70	0.75

superiority of ECP over EQC is much more pronounced in long-duration charging instances, which we omit for brevity.

We next compare ECP with RH-LP and RH-ECP on the small-scale (six-period) numerical instances in Table 9. We find that RH-ECP performs the best among all the policies in terms of both the mean and standard deviation of the total cost. When *C* is large, even the (static) ECP performs better than the sensible adaptive policy RH-LP on many instances. This suggests our proposed ECP approach is sufficiently adaptive to customer types (arrival/departure time and charging requirement) and time period.

7.4.4. Role of entropic dominance constraints. We now examine the role of entropic dominance constraints in constructing our ECP formulation. Recall in §4.2, we obtain our ECP by leveraging the DRO framework and use the ambiguity set  $\mathcal{F}$ , which is the intersection of  $\mathcal{F}^1$ , containing the entropic dominance constraints, and  $\mathcal{F}^2$ , a commonly used ambiguity set with the support and mean information. To examine the role of  $\mathcal{F}^1$ , we compare the DRO formulation with  $\mathcal{F}^1$ , which is DRO-Ent (in turn approximated by ECP-C as seen in §4.2), with another DRO formulation without  $\mathcal{F}^1$  as follows:

$$
\min_{\boldsymbol{x} \in \mathcal{X}} \sup_{\mathbb{P} \in \mathcal{F}^2} \mathbb{E}_{\mathbb{P}} \left[ c(\boldsymbol{x}, \tilde{\boldsymbol{z}}) \right]. \tag{No-Ent}
$$

In other words, No-Ent is the same as DRO-Ent except that the ambiguity set is  $\mathcal{F}^2$  instead of  $\mathcal{F}$ . Hence, we need to compare ECP-C with No-Ent, which can be reformulated as a tractable LP.

$\overline{C}$	Method	Objective	Total cost			
			Mean	Std		
30	ECP-C	899.21	805.00	67.55		
	No-Ent	1366.00	834.90	71.43		
25	ECP-C	891.43	783.66	57.13		
	$No\text{-}Ent$	1227.68	818.03	60.40		
20	ECP-C	854.78	717.73	44.42		
	No-Ent	1082.74	756.78	46.25		
15	ECP-C	779.44	600.40	35.20		
	No-Ent	916.34	639.09	36.35		
10	ECP-C	639.18	441.34	27.71		
	$No\text{-}Ent$	701.80	478.87	29.32		
5	ECP-C	411.01	250.43	19.43		
	No-Ent	421.15	282.86	19.72		

Table 10 Objective values and performances of ECP-C and No-Ent

Table 10 compares the objective values and performance evaluations of ECP-C and No-Ent given different values of the number of chargers  $C$ . We see that as  $C$  decreases, the objective value of ECP-C tends to coincide with that of No-Ent. For instance, when *C* = 5, their objective values are very close, differing by about  $2\%$ . Nevertheless, when the solutions from ECP-C and No-Ent are evaluated on 10,000 sample paths, the mean total cost of ECP-C is about 11% smaller than that of No-Ent. This observation holds for all values of *C* considered in the table. This highlights the important role entropic dominance constraints play in the ambiguity set in constructing ECP.

#### 7.5. Managerial Insights in EV Charging Scheduling

We use ECP-C to first study the value of customer desired departure time information and then examine the effect of the composition of demand charge on electricity load.

7.5.1. Value of desired departure time information. We study this value by comparing the performance of ECP-C with that of the maximum-speed charging policy, which charges vehicles using maximal charging speed (power capacity) until completion on a first-come-first-served basis and is often used in practice and literature (Zhang et al. 2021). Denote by  $x^M$  the charging decision under this maximum-speed charging policy. Then  $x_{v,t}^M = \max\{0, \min\{K, u_v - (t - s_v) \cdot K\}\}\)$  for all  $v \in [V]$  and  $t \in \mathcal{T}_v$ . We evaluate the maximum-speed charging policy in terms of both the mean and standard deviation of the total cost analogous to the evaluation of ECP-C. We also evaluate under both policies the mean of the maximum electricity load during the entire horizon, on-peak hours, mid-peak hours, and off-peak hours. Table 11 lists all the performance measures of ECP-C relative to those of the maximum-speed charging policy. Using column 2 as an example, if we denote  $\pi^M = \mathbb{E}_{\mathbb{P}^C}[c(\mathbf{x}^M, \tilde{\mathbf{z}})]$ , column 2 represents  $(\pi^M - \pi^E)/\pi^M \times 100\%$ .

$\overline{C}$	Total cost		Expected maximum load			
	Mean	Std			Overall $\vert$ On-peak $\vert$ Mid-peak $\vert$ Off-peak	
15 <sup>1</sup>			$13.60\%$   27.87\%   23.64\%	23.56%	21.43%	10.33%
20		$12.12\%$   25.49\%   20.99\%		20.68%	18.56%	8.41\%
25		$11.22\%$   $21.80\%$   $18.91\%$		19.00%	16.03%	8.70%
30		$10.99\%$   $19.02\%$   $17.94\%$		18.89%	13.65%	8.59%

Table 11 Performance improvement of ECP-C over the maximum-speed charging policy

*Notes.* The 2nd and 3rd columns are the percentage reduction in the mean and standard deviation, respectively, of the total cost of ECP-C compared with the maximum-speed charging policy. Analogously, the 4th to 7th columns are the percentage difference of the mean of the maximum load over the entire time horizon, on-peak hours, mid-peak hours, and off-peak hours, respectively. Note also that the expected maximum load over the entire horizon (4th column) differs from the maximum of the expected maximum load during different time intervals (in the 5th to 7th columns), as the maximum load may occur in different time intervals.

As seen in Table 11, compared with the maximum-speed charging policy, ECP-C results in a lower mean and standard deviation of the total cost, e.g., with 30 chargers, 10*.*99% and 19.02%, respectively. This lower cost of ECP-C is reflected in the lower expected maximum electricity load during any time interval, implying a more smooth load. Moreover, the lower expected maximum load from ECP-C alleviates the potential overcharge to EVs under the maximum-speed charging policy, and thus ECP-C can also lower the risk of EV battery degradation (Thompson 2018). When station capacity decreases, the difference between ECP-C and the maximum-speed charging policy increases. This difference can be regarded as the value of information on the customer departure time, as ECP-C differs from the maximum-speed charging policy in that while the former uses this information, the latter does not. This value is even higher for the long-duration charging instances, which we omit for brevity. These results suggest that a charging station can benefit significantly from incorporating the customer desired departure information, as our SP model does, to smooth the electricity load and reduce the total cost.

7.5.2. Effect of the composition of the total demand charge on electricity load. The rate structure of most utility firms has three common components of Time-Of-Use (TOU) demand charges, namely on-peak, mid-peak, and all-period demand changes (see their corresponding hours in Figure 5). We examine the following compositions of the total demand charge: (i) without any demand charge; (ii) with only on-peak demand charge; (iii) with only mid-peak demand charge; (iv) with both on-peak and mid-peak demand charges; (v) with only all-period demand charge; and (vi) with both all-period and on-peak demand charges. We compare these with the baseline case where all three components exist. Given each composition of the total demand charge, we compute  $x<sup>E</sup>$ , the optimal solution to ECP-C, and then use it to compute the mean and standard deviation of the electricity load in the charging station for each period, which we plot in Figure 5. We make the following observations:

*Observation 1. In the absence of any demand charge, the TOU energy charge creates an artificial electricity load spike around the transitions between any two time intervals.* As seen in Figure 5a, the expected load swings significantly around the hours between any two sets of time intervals: around hour 9 (between off-peak and mid-peak hours) and around hours 12 and 18 (between midpeak and on-peak hours). This means that the TOU energy charge (see tariff structure in  $\S7.1$ ) alone is not sufficient to smooth the electricity load and may even create artificial spikes, putting undue stress on the electric grid. This is consistent with the observation in a similar setting in practice (Zhang and Qian 2018): At midnight, there is a huge spike in electricity charging load from residential customers, because their electricity tariff structures do not include any demand charge while their electricity price at midnight is lower than that at hour 11 pm. We show this observation holds when the decisions are near optimal in the EV charging management setting.

*Observation 2. Either an on-peak or mid-peak demand charge alone cannot smooth the electricity load and may even create a larger load swing than the case without any demand charge.* As shown in Figure 5b, when only an on-peak demand charge exists, the electricity load is smoothed only within on-peak hours, but not over time. In particular, there are some artificial spikes in the load around all transition hours (i.e., between on-peak and mid-peak hours and between mid-peak and off-peak hours) because the charging station simply shifts the high load from on-peak hours to other hours.

Figure 5 (Color Online) Electricity load under ECP-C for different compositions of the demand charge given  $C = 30$ 



*Notes.* The black, white, and gray regions on the x-axis represent on-peak hours (hours 13 to 18), mid-peak hours (hours 9 to 12 and hours 19 to 23), and off-peak hours (hours 1 to 8 and hour 24), respectively.

These spikes are even higher than those within the same hours when there is no demand charge, and thus impose more stress on the electric grid. When there is only a mid-peak demand charge, an analogous pattern occurs (the corresponding figure is omitted for brevity). Similarly, when both an on-peak demand charge and mid-peak demand charge exist, the load is smoothed only during these combined hours, and load spikes occur around hours between mid-peak and off-peak hours (figure omitted for brevity).

*Observation 3. An all-period demand charge alone is sufficient to smooth electricity load over time.* In contrast to the case in which there is no demand charge or only either an on-peak or mid-peak demand charge, an all-period demand charge results in a smooth electricity load (see Figure 5c). The load curve for this case almost coincides with that for the case in which all three components of demand charge exist (see Figure 5d). This is because, with an all-period demand charge alone, the charging station is incentivized to reduce the highest load over the entire time horizon.

#### 7.6. Joint Pricing and Scheduling

We use both ECP and SAA approaches to solve the joint pricing and scheduling problem in  $(17)$ . For simplicity, we set  $C = 30$ ; we use only all-period demand charge and set  $d = \frac{60.930}{kW}$ ; we set  $[p,\overline{p}]$  as  $[0.3, 0.7]$ ,  $r_v = 4/3$  for all  $v \in [V]$ , and  $\overline{\lambda}_v$  so that the resulting arrival rates when  $p = 0.5$ (the mid point of  $[p,\overline{p}]$ ) are the same as those in §7.1. We set the stopping tolerance  $\delta = 0.01$ .

Figure 6 (Color Online) Performance of the alternating optimization algorithm on JPS-C



*Notes.* The objective values on the dotted line at integer number of iterations correspond to the optimal values of the scheduling optimization, while those at the half-integer number of iterations correspond to the optimal values of the pricing optimization. The asterisks  $(*)$  correspond to the expected profit in (17) evaluated at JPS-C solutions after each iteration.

We solve  $(17)$  using JPS-C with an alternating optimization algorithm analogous to Algorithm 1. Figure 6 plots the objective values of JPS-C and evaluations of JPS-C (i.e., substituting the JPS-C solutions to  $(17)$  on 10,000 sample paths) at two different initial values of  $p$ , 0*.*3 and 0*.7*. We find that the algorithm converges in four iterations to give the same price  $p = 0.49$ . Hence, our method is fast and not sensitive to the initial price *p*. Although our algorithm is not guaranteed to converge to a globally optimal solution of JPS-C, it performs well on our numerical instances, where  $p = 0.49$ can be numerically verified as the globally optimal price.

We also solve (17) using SAA, for which we cannot directly extend (14) as the arrival rates of Poisson random variables,  $\lambda$ , depend on the pricing decision, p. Thus, we first enumerate p and solve for each value of *p* the corresponding SAA problem to obtain an expected profit, and finally select the value of *p* which results in the highest expected profits. We use 8,000 samples and enumerate *p* from 0.3 to 0.7 with a stepsize of 0.005. Figure 7 plots the expected profits under SAA at different  $p \in [0.45, 0.54]$  (while the rest of the range is omitted), which shows that the optimal *p* under SAA is 0.48. (Note that JPS-C does not need to enumerate *p* and instead uses the alternating algorithm mentioned above, so there is only one objective value under JPS-C in Figure 7.)



#### Figure 7 (Color Online) Expected profits under SAA and JPS-C

We evaluate SAA by substituting its solutions into  $(17)$  on 10,000 samples to obtain the corresponding expected profits. The optimal price  $p = 0.48$  under SAA leads to an evaluation of the expected profits under SAA to be 739.62, lower than the expected profits under JPS-C, which is 749.91. The difference between these two is significant as the standard error of the expected profits under both is 0.72, which is very low compared to the difference between their expected profits. The superiority of JPS-C is due to ECP-C's better performance in scheduling optimization than SAA, as mentioned in §7.3. Moreover, JPS-C runs much faster, while SAA requires solving 81 large-scale SAA instances, each of which can be time-consuming. In sum, ECP (or JPS-C) is a better approach than SAA in solving the joint pricing and scheduling problem.

#### 8. Conclusions and Future Work

We use an ECP approach to examine the EV charging management of a service provider who faces uncertain arrivals of customers with heterogeneous arrival/desired departure times and charging requirements. We formulate this problem as an SP where customer arrivals follow a joint Poisson distribution to minimize the expected total electricity cost which includes demand charges. As the SP is large-scale, we develop ECP approximations. When the service provider is uncapacitated, we develop an ECP approximation by bounding the largest order statistic, introduced by the demand charge, using exponential cones based on the epigraph of MGFs of the Poisson customer arrivals. We show that this ECP gives an upper bound of the SP and characterize a bound on the gap between their theoretical performances. In the capacitated case, we develop another ECP approximation by also using the idea from DRO of employing an entropic dominance ambiguity set. Rather than mitigating distributional ambiguity, we use DRO to derive an ECP as a tractable upper bound of the corresponding SP. Moreover, we extend our ECP approach to jointly optimize pricing and scheduling and propose an alternating optimization algorithm to solve the resulting ECP approximations efficiently.

We then benchmark our ECP approach with two common approaches in solving SPs, SAA and DRO, on numerical instances calibrated to real tariff and EV customer arrival data. As our numerical instances are large-scale (with about 80,000 random variables and 700,000 decision variables), we show that DRO cannot be solved on these instances; ECP runs about 50 times faster than SAA and results in not only a lower standard deviation of the total cost but also a lower mean of this cost because SAA does not scale well and cannot be solved to optimality in a reasonable amount of time. We also find the optimality gap of ECP is at most around 4% on all instances. The superior performance of ECP over SAA continues to hold when considering three practical implementation issues: a finer time discretization, a data-driven setting when there are estimation errors in the arrival rate (due to insufficient data) or even the mis-specification of the arrival distribution, and an adaptive charging setting where uncertainty is revealed over time. In the extended model of joint pricing and scheduling, ECP results in a higher expected profit than SAA and also runs faster. In addition, ECP is a deterministic approximation, while SAA depends on samples generated. Therefore, ECP emerges as the best approach in solving the EV charging problem.

We finally use our ECP approach to generate insights on a similar set of data-calibrated numerical instances and find the following: First, it is beneficial for the charging service provider to use the customer departure information, as the ECP approach outperforms by about 11% a commonly-used maximum-speed charging policy which ignores such information. Second, a TOU energy charge alone (i.e., without any demand charge) creates artificial spikes in the electricity load, consistent with the observation in practice, which we show to hold when the decisions are near optimal in the EV management setting. While either an on-peak or mid-peak demand charge shifts the artificial spikes and could not smooth the electricity load, an all-period demand charge alone can result in a load as smooth over time as the case in which all three types of demand charge exist.

Our ECP approach is appealing for use in practice by a charging service provider as it outperforms various benchmarks considering different practical implementation issues. The managerial insights can give guidance for charging service providers in terms of the value of collecting departure time information from customers as well as policymakers in designing the electricity tariff structure to smooth electricity load.

There are a few promising avenues to extend this work. First, our SP model assumes that the departure time provided by customers is accurate. In practice, customers may underreport their departure time to reduce the risk of insufficient charge upon pickup, and thus it is potentially beneficial for the service provider to charge less aggressively to minimize the total expected cost. Considering this behavior should prove to be fruitful. Second, our methods to construct ECP approximations may be applicable to other two-stage stochastic linear programs with fixed recourse. Acknowledgement. The authors thank the entire review team for the wonderful suggestions that have helped significantly improve the paper. The first author gratefully acknowledges the financial support from the Singapore Ministry of Education with the 2019 Academic Research Fund Tier 3 [Grant MOE-2019-T3-1-010].

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## Online Appendix for *An Exponential Cone Programming Approach for Managing Electric Vehicle Charging*

#### Appendix A: Proofs.

**Proof of Theorem 1.** For simplicity, we denote  $h(x) \triangleq \mathbb{E}_{\mathbb{P}^\infty} [c(x, \tilde{z})]$ ; for any given  $\mu > 0$ , let

$$
\bar{h}(\boldsymbol{x};\mu) \triangleq \sum_{s \in [T]} e_s f_s(\boldsymbol{x}, \boldsymbol{\lambda}) + d \left( \mu \ln \sum_{t \in [T]} \exp \left( \sum_{v \in \mathcal{V}_t} \lambda_v \phi(x_{v,t}/\mu) \right) + \max_{t \in [T]} f_t(\boldsymbol{x}, \boldsymbol{\lambda}) \right);
$$
  

$$
\underline{h}(\boldsymbol{x}) \triangleq \sum_{s \in [T]} e_s f_s(\boldsymbol{x}, \boldsymbol{\lambda}) + d \left( \max_{t \in [T]} f_t(\boldsymbol{x}, \boldsymbol{\lambda}) \right).
$$

Let  $(\bar{x}, \bar{\mu})$  be the optimal solution to  $\inf_{x \in \mathcal{X}, \mu > 0} \bar{h}(x; \mu)$ , or equivalently the optimal solution to ECP-U by (6) in the proof of Proposition 1, and  $\mu^*$  be the optimal solution to  $\inf_{\mu>0} \bar{h}(x^*;\mu)$ . Note that  $h(x)$  is a lower bound of  $h(x)$  for any  $x \in \mathcal{X}$  because  $\max_{t \in [T]} f_t(x, z)$  is convex in z. So due to Jensen's inequality, we have

$$
\mathbb{E}_{\mathbb{P}^\infty}\left[\max_{t\in[T]}f_t(\boldsymbol{x}, \tilde{\boldsymbol{z}})\right] \geq \max_{t\in[T]}f_t(\boldsymbol{x}, \mathbb{E}_{\mathbb{P}^\infty}\left[\tilde{\boldsymbol{z}}\right]) = \max_{t\in[T]}f_t(\boldsymbol{x}, \boldsymbol{\lambda}).
$$

And  $\bar{h}(\mathbf{x};\mu)$  is an upper bound of  $h(\mathbf{x})$  from Proposition 1 for any  $\mu > 0$ . Based on the lower bound and upper bound of  $\bar{x}$ , we bound the performance of  $\bar{x}$  in the following:

$$
h(\bar{\mathbf{x}}) - h(\mathbf{x}^*) = (h(\bar{\mathbf{x}}) - \bar{h}(\bar{\mathbf{x}}; \bar{\mu})) + (\bar{h}(\bar{\mathbf{x}}; \bar{\mu}) - \bar{h}(\mathbf{x}^*; \mu^*)) + (\bar{h}(\mathbf{x}^*; \mu^*) - h(\mathbf{x}^*))
$$
  
\n
$$
\leq \bar{h}(\mathbf{x}^*; \mu^*) - h(\mathbf{x}^*)
$$
  
\n
$$
\leq \bar{h}(\mathbf{x}^*; \mu^*) - \underline{h}(\mathbf{x}^*)
$$
  
\n
$$
= d\mu^* \ln \sum_{t \in [T]} \exp\left(\sum_{v \in \mathcal{V}_t} \lambda_v \phi(x_{v,t}^*/\mu^*)\right)
$$
  
\n
$$
= d \inf_{\mu > 0} \left(\mu \ln \sum_{t \in [T]} \exp\left(\sum_{v \in \mathcal{V}_t} \lambda_v \phi(x_{v,t}^*/\mu)\right)\right),
$$
\n(24)

where in the first inequality we have  $h(\bar{x}) \leq \bar{h}(\bar{x}; \bar{\mu})$  and  $\bar{h}(\bar{x}; \bar{\mu}) \leq \bar{h}(\bar{x}^*, \mu^*)$  due to definition of  $(\bar{x}, \bar{\mu})$  and  $\bar{h}(\mathbf{x}; \mu)$ . Note that for any  $\mathbf{x} \in \mathcal{X}$  and  $\mu > 0$ , we have

$$
\mu \ln \sum_{t \in [T]} \exp \left( \sum_{v \in \mathcal{V}_t} \lambda_v \phi(x_{v,t}^*/\mu) \right) \le \mu \ln \sum_{t \in [T]} \exp \left( \sum_{v \in \mathcal{V}_t} \lambda_v \phi \left( \frac{\min\{u_v, K\}}{\mu \eta} \right) \right)
$$

as  $x_{v,t}^* \le \min\{u_v, K\}/\eta$  and  $\mu \ln \sum_{t \in [T]} \exp\left(\sum_{v \in \mathcal{V}_t} \lambda_v \phi(\cdot/\mu)\right)$  is increasing. Thus, by taking the minimum over  $\mu > 0$ , we obtain the performance guarantee for  $\bar{x}$  in (12).

**Proof of Lemma 1.**  $\mathcal{F}^1 \cap \mathcal{F}^2 \subseteq \mathcal{F}$  is clearly true because  $\mathcal{F}^1 \cap \mathcal{F}^2$  includes all the constraints in  $\mathcal{F}$ . To prove  $\mathcal{F} \subseteq \mathcal{F}^1 \cap \mathcal{F}^2$ , we next show that any  $\tilde{z}$  that satisfies the following also satisfies  $\mathbb{E}_{\mathbb{P}}[\tilde{z}] \leq \lambda$ :

$$
\ln \mathbb{E}_{\mathbb{P}} \left[ \exp \left( \boldsymbol{\theta}^{\prime} \tilde{\boldsymbol{z}} \right) \right] \leq \sum_{v \in [V]} \lambda_{v} \left( e^{\theta_{v}} - 1 \right) \ \forall \boldsymbol{\theta} \geq \mathbf{0}.
$$

For each  $v \in [V]$ , consider  $\theta_v \triangleq (0, ..., \theta_v, 0, ..., 0)$ , i.e., every component is 0 except one, which is strictly positive. Then we have  $\ln \mathbb{E}_{\mathbb{P}} [\exp(\theta' \tilde{z})] = \ln \mathbb{E}_{\mathbb{P}} [\exp(\theta_v \tilde{z}_v)] \leq \lambda_v (e^{\theta_v} - 1)$ . Taking exponential function on both

sides we have an equivalent inequality  $\mathbb{E}_{\mathbb{P}}[\exp(\theta_v \tilde{z}_v)] \leq \exp(\lambda_v(e^{\theta_v}-1))$ . Using Taylor's expansion on both sides of the inequality, we have for each  $\theta_v > 0$ 

$$
\mathbb{E}_{\mathbb{P}}\left[\exp\left(\theta_{v}\tilde{z}_{v}\right)\right] = \mathbb{E}_{\mathbb{P}}\left[1 + \theta_{v}\tilde{z}_{v} + \sum_{k=2}^{\infty}\frac{(\theta_{v}\tilde{z}_{v})^{k}}{k!}\right] = 1 + \mathbb{E}_{\mathbb{P}}\left[\theta_{v}\tilde{z}_{v}\right] + \sum_{k=2}^{\infty}\mathbb{E}_{\mathbb{P}}\left[\frac{(\theta_{v}\tilde{z}_{v})^{k}}{k!}\right]
$$

$$
= 1 + \mathbb{E}_{\mathbb{P}}\left[\theta_{v}\tilde{z}_{v}\right] + o(\theta_{v}) \le \exp\left(\lambda_{v}(e^{\theta_{v}} - 1)\right) = 1 + \lambda_{v}\theta_{v} + o(\theta_{v}),
$$

where the second equality is from Fubini's theorem. After simplifying and dividing  $\theta_v$  on both sides of

$$
\mathbb{E}_{\mathbb{P}}\left[\theta_v\tilde{z}_v\right] + o(\theta_v) \leq \lambda_v\theta_v + o(\theta_v),
$$

then letting  $\theta_v \to 0^+$ , we get  $\mathbb{E}_{\mathbb{P}}[\tilde{z}_v] \leq \lambda_v$  for each  $v \in [V]$ . Hence  $\mathbb{E}_{\mathbb{P}}[\tilde{z}] \leq \lambda$ .

**Proof.** of Proposition 2. The support constraint of  $\mathbb{P}^C$  in (1) are in the third line in the definition of *F*. Let  $\tilde{z} \sim \mathbb{P}^C$  and  $\tilde{w} \sim \mathbb{P}^{\infty}$  (i.e., the vector of independent Poisson random variables  $\tilde{w}_v$  with arrival rate  $\lambda_v > 0$ , and denote the joint distribution of  $(\tilde{z}, \tilde{w})$  as  $\overline{\mathbb{P}}$ . By the definition of  $\tilde{z}$ , we know  $\tilde{z} \leq \tilde{w}$  holds almost surely. Then  $\theta' \tilde{z} \leq \theta' \tilde{w}$  holds for all  $\theta \geq 0$  with probability one. So we have

$$
\ln \mathbb{E}_{\mathbb{P}^C} \left[ \exp \left( \boldsymbol{\theta}^{\prime} \tilde{\boldsymbol{z}} \right) \right] = \ln \mathbb{E}_{\bar{\mathbb{P}}} \left[ \exp \left( \boldsymbol{\theta}^{\prime} \tilde{\boldsymbol{z}} \right) \right] \leq \ln \mathbb{E}_{\bar{\mathbb{P}}} \left[ \exp \left( \boldsymbol{\theta}^{\prime} \tilde{\boldsymbol{w}} \right) \right] = \ln \mathbb{E}_{\mathbb{P}^{\infty}} \left[ \exp \left( \boldsymbol{\theta}^{\prime} \tilde{\boldsymbol{w}} \right) \right] = \sum_{v \in [V]} \lambda_v (e^{\theta_v} - 1),
$$

where the inequality is due to the monotonicity of  $\ln \mathbb{E}_{\bar{p}} [\exp(\cdot)]$  and the third equality follows from (4). Hence  $\mathbb{P}^C \in \mathcal{F}$ .

To prove Proposition 3, we establish the next lemma first.

LEMMA 3. For a piece-wise linear convex function  $g(z) \triangleq \max_{i \in [I]} \{x' A_i z\}$  and any random vector  $\tilde{z} \sim \mathbb{P}$ , *we have*

$$
\mathbb{E}_{\mathbb{P}}\left[g(\tilde{\boldsymbol{z}})\right] = \inf_{\boldsymbol{y}} \mathbb{E}_{\mathbb{P}}\left[\max_{i \in [I]} \left\{\boldsymbol{y}'\boldsymbol{A}_i \tilde{\boldsymbol{z}}\right\}\right] + \mathbb{E}_{\mathbb{P}}\left[\max_{i \in [I]} \left\{\left(\boldsymbol{x} - \boldsymbol{y}\right)' \boldsymbol{A}_i \tilde{\boldsymbol{z}}\right\}\right].
$$
\n(25)

*Proof of Lemma 3.* Note (25) holds when  $y = 0$  for all  $i \in [I]$ , we have

$$
\mathbb{E}_{\mathbb{P}}\left[g(\tilde{\bm{z}})\right] \geq \inf_{\bm{y}} \mathbb{E}_{\mathbb{P}}\left[\max_{i \in [I]} \left\{\bm{y'A}_i \tilde{\bm{z}}\right\}\right] + \mathbb{E}_{\mathbb{P}}\left[\max_{i \in [I]} \left\{(\bm{x}-\bm{y})'\bm{A}_i \tilde{\bm{z}}\right\}\right].
$$

Note that for all  $\bm{z}$  and  $\bm{y}$ , we have  $g(\bm{z}) \leq \max_{i \in [I]} \{\bm{y'}\bm{A}_i\bm{z}\} + \max_{i \in [I]} \{(\bm{x} - \bm{y})'\bm{A}_i\bm{z}\}$ . Hence for all  $\bm{y}$ , we have  $\mathbb{E}_{\mathbb{P}}[g(\tilde{\boldsymbol{z}})] \leq \mathbb{E}_{\mathbb{P}}[\max_{i \in [I]} \{\boldsymbol{y}' \boldsymbol{A}_i \tilde{\boldsymbol{z}}\} + \mathbb{E}_{\mathbb{P}}[\max_{i \in [I]} \{(\boldsymbol{x} - \boldsymbol{y})' \boldsymbol{A}_i \tilde{\boldsymbol{z}}\}].$  By taking infimum over  $\boldsymbol{y}$ , we get the conclusion.  $\Box$ 

Proof of Proposition 3. Since DRO-Ent is an upper bound of  $(3)$ , we then obtain an upper bound of DRO-Ent, which then becomes an upper bound of (3). We can then bound  $\sup_{\mathbb{P}\in\mathcal{F}}\mathbb{E}_{\mathbb{P}}[c(\mathbf{x},\tilde{\mathbf{z}})]$  as follow,

$$
\sup_{\mathbb{P}\in\mathcal{F}}\mathbb{E}_{\mathbb{P}}\left[c(\boldsymbol{x},\tilde{\boldsymbol{z}})\right]\leq \sup_{\mathbb{P}\in\mathcal{F}}\mathbb{E}_{\mathbb{P}}\left[\sum_{s\in[T]}e_{s}f_{s}(\boldsymbol{x},\tilde{\boldsymbol{z}})\right]+d\sup_{\mathbb{P}\in\mathcal{F}}\mathbb{E}_{\mathbb{P}}\left[\max_{t\in[T]} \{f_{t}(\boldsymbol{x},\tilde{\boldsymbol{z}})\}\right].
$$
\n(26)

Then we derive an upper bound for each of the two terms on the RHS of (26). For the first term, due to the fact  $\mathcal{F} \subset \mathcal{F}^2$  (based on Lemma 1), we have

$$
\sup_{\mathbb{P}\in\mathcal{F}}\mathbb{E}_{\mathbb{P}}\left[\sum_{s\in[T]}e_s f_s(\boldsymbol{x}, \tilde{\boldsymbol{z}})\right]\leq \sup_{\mathbb{P}\in\mathcal{F}^2}\mathbb{E}_{\mathbb{P}}\left[\sum_{s\in[T]}e_s f_s(\boldsymbol{x}, \tilde{\boldsymbol{z}})\right].
$$

Using standard duality results, we have

$$
\sup_{\mathbb{P}\in\mathcal{F}^2} \mathbb{E}_{\mathbb{P}}\left[\sum_{s\in[T]} e_s f_s(\boldsymbol{x}, \tilde{\boldsymbol{z}})\right] = \inf_{a,b\geq 0} a + b'\lambda
$$
\n
$$
\text{s.t. } \sum_{s\in[T]} e_s f_s(\boldsymbol{x}, \boldsymbol{z}) \leq a + b'\boldsymbol{z} \qquad \forall \boldsymbol{z} \in \mathcal{Z}
$$
\n
$$
= \inf_{a,b\geq 0} a + b'\lambda
$$
\n
$$
\text{s.t. } \sup_{\boldsymbol{z}\in\mathcal{Z}} \sum_{s\in[T]} e_s f_s(\boldsymbol{x}, \boldsymbol{z}) - b'\boldsymbol{z} \leq a
$$
\n
$$
(27)
$$

Note that the dual of  $\sup_{z\geq 0} \left\{ \sum_{s\in[T]} e_s \sum_{v\in\mathcal{V}_s} x_{v,s} z_v - \sum_{v\in[V]} b_v z_v \mid \sum_{v\in\mathcal{V}_t} z_v \leq C, \forall t\in[T] \right\} \leq a$  is

$$
\inf_{\nu \geq 0} \left\{ C \sum_{t \in [T]} \nu_t \mid \sum_{s \in \mathcal{T}_v} x_{v,s} e_s - b_v \leq \sum_{t \in \mathcal{T}_v} \nu_t, \forall v \in [V] \right\} \leq a,
$$

which is equivalent to equation (13g) and (13h). Therefore, (27) becomes

$$
\sup_{\mathbb{P}\in\mathcal{F}^2} \mathbb{E}_{\mathbb{P}}\left[\sum_{s\in[T]} e_s f_s(\boldsymbol{x}, \tilde{\boldsymbol{z}})\right] = \inf_{\boldsymbol{b}\geq 0, \boldsymbol{\nu}\geq 0} a + \boldsymbol{b}'\boldsymbol{\lambda}
$$
\n
$$
\text{s.t.} \quad (13\text{g}), (13\text{h}). \tag{28}
$$

We next obtain an upper bound of  $\sup_{\mathbb{P}\in\mathcal{F}} \mathbb{E}_{\mathbb{P}} \left[\max_{t\in[T]}\{f_t(\boldsymbol{x}, \tilde{\boldsymbol{z}})\}\right]$ . Since  $c(\boldsymbol{x}, \boldsymbol{z})$  is a piece-wise linear convex function of *z* given *x*, we have

$$
\sup_{\mathbb{P}\in\mathcal{F}}\mathbb{E}_{\mathbb{P}}\left[\max_{t\in[T]}\{f_t(\boldsymbol{x},\tilde{\boldsymbol{z}})\}\right] = \sup_{\mathbb{P}\in\mathcal{F}}\inf_{\boldsymbol{y}}\left\{\mathbb{E}_{\mathbb{P}}\left[\max_{t\in[T]}\{f_t(\boldsymbol{y},\tilde{\boldsymbol{z}})\}\right] + \mathbb{E}_{\mathbb{P}}\left[\max_{t\in[T]}\{f_t(\boldsymbol{x}-\boldsymbol{y},\tilde{\boldsymbol{z}})\}\right]\right\}
$$
\n
$$
\leq \inf_{\boldsymbol{y}}\sup_{\mathbb{P}\in\mathcal{F}}\left\{\mathbb{E}_{\mathbb{P}}\left[\max_{t\in[T]}\{f_t(\boldsymbol{y},\tilde{\boldsymbol{z}})\}\right] + \mathbb{E}_{\mathbb{P}}\left[\max_{t\in[T]}\{f_t(\boldsymbol{x}-\boldsymbol{y},\tilde{\boldsymbol{z}})\}\right]\right\}
$$
\n
$$
\leq \inf_{\boldsymbol{y}}\left\{\sup_{\mathbb{P}\in\mathcal{F}^1}\mathbb{E}_{\mathbb{P}}\left[\max_{t\in[T]}\{f_t(\boldsymbol{y},\tilde{\boldsymbol{z}})\}\right] + \sup_{\mathbb{P}\in\mathcal{F}^2}\mathbb{E}_{\mathbb{P}}\left[\max_{t\in[T]}\{f_t(\boldsymbol{x}-\boldsymbol{y},\tilde{\boldsymbol{z}})\}\right]\right\}
$$
\n
$$
\leq \inf_{\boldsymbol{y}\geq 0}\left\{\sup_{\mathbb{P}\in\mathcal{F}^1}\mathbb{E}_{\mathbb{P}}\left[\max_{t\in[T]}\{f_t(\boldsymbol{y},\tilde{\boldsymbol{z}})\}\right] + \sup_{\mathbb{P}\in\mathcal{F}^2}\mathbb{E}_{\mathbb{P}}\left[\max_{t\in[T]}\{f_t(\boldsymbol{x}-\boldsymbol{y},\tilde{\boldsymbol{z}})\}\right]\right\},
$$
\n(29)

where the equality is from Lemma 3 and the second inequality is from Lemma 1. We next bound  $\sup_{\mathbb{P}\in\mathcal{F}^1}\mathbb{E}_{\mathbb{P}}\left[\max_{t\in[T]}f_t(\boldsymbol{x},\tilde{\boldsymbol{z}})\right]$  and  $\sup_{\mathbb{P}\in\mathcal{F}^2}\mathbb{E}_{\mathbb{P}}\left[\max_{t\in[T]}f_t(\boldsymbol{x},\tilde{\boldsymbol{z}})\right]$ . Similar to the proof of Proposition 1, for any  $y \geq 0$ , we have

$$
\begin{aligned} \sup_{\mathbb{P}\in\mathcal{F}^1}\mathbb{E}_{\mathbb{P}}\left[\max_{t\in[T]}f_t(\bm{y},\tilde{\bm{z}})\right] &=\sup_{\mathbb{P}\in\mathcal{F}^1}\mathbb{E}_{\mathbb{P}}\left[\max_{t\in[T]}\left\{f_t(\bm{y},\tilde{\bm{z}})-f_t(\bm{y},\bm{\lambda})+f_t(\bm{y},\bm{\lambda})\right\}\right] \\ &\leq \sup_{\mathbb{P}\in\mathcal{F}^1}\mathbb{E}_{\mathbb{P}}\left[\max_{t\in[T]}\left\{f_t(\bm{y},\tilde{\bm{z}})-f_t(\bm{y},\bm{\lambda})\right\}\right]+\max_{t\in[T]}f_t(\bm{y},\bm{\lambda}) \\ &\leq \mu\ln\sum_{t\in[T]}\sup_{\mathbb{P}\in\mathcal{F}^1}\mathbb{E}_{\mathbb{P}}\left[\exp\left((f_t(\bm{y},\tilde{\bm{z}})-f_t(\bm{y},\bm{\lambda}))/\mu\right)\right]+\max_{t\in[T]}f_t(\bm{y},\bm{\lambda}) \\ &\leq \mu\ln\sum_{t\in[T]}\exp\left(\sum_{v\in\mathcal{V}_t}\lambda_v\left(e^{y_{v,t}/\mu}-1-y_{v,t}/\mu\right)\right)+\max_{t\in[T]}f_t(\bm{y},\bm{\lambda}), \end{aligned}
$$

which has the same form as (7). By the same argument as in the proof of Proposition 1, we can bound  $\sup_{\mathbb{P}\in\mathcal{F}^1} \mathbb{E}_{\mathbb{P}}\left[\max_{t\in[T]} f_t(\mathbf{y}, \tilde{\mathbf{z}})\right]$  by

$$
\inf_{\kappa,\gamma,\mu>0,\xi,\zeta} d(\kappa+\gamma)
$$
  
s.t. (13a) – (13d). (30)

Similar to (28), we bound  $\sup_{\mathbb{P}\in\mathcal{F}^2} \mathbb{E}_{\mathbb{P}} \left[ \max_{t\in[T]} \{ f_t(\mathbf{x}-\mathbf{y}, \tilde{\mathbf{z}}) \} \right]$  using the strong duality results, i.e.,

$$
\sup_{\mathbb{P}\in\mathcal{F}^2} \mathbb{E}_{\mathbb{P}}\left[\max_{t\in[T]} \{f_t(\mathbf{x}-\mathbf{y}, \tilde{\mathbf{z}})\}\right] = \inf_{\alpha, \beta \ge 0} \alpha + \beta' \lambda
$$
\n
$$
\text{s.t.} \quad f_t(\mathbf{x}-\mathbf{y}, \mathbf{z}) \le \alpha + \beta' \mathbf{z} \quad \forall \mathbf{z} \in \mathcal{Z}, t \in [T]
$$
\n
$$
= \inf_{\alpha, \beta \ge 0} \alpha + \beta' \lambda
$$
\n
$$
\text{s.t.} \quad \sup_{\mathbf{z} \in \mathcal{Z}} f_t(\mathbf{x}-\mathbf{y}, \mathbf{z}) - \beta' \mathbf{z} \le \alpha \ \forall t \in [T]
$$
\n
$$
= \inf_{\alpha, \rho \ge 0, \beta \ge 0} \alpha + \beta' \lambda
$$
\n
$$
\text{s.t.} \quad (13e), (13f)
$$
\n
$$
-\beta_v \le \sum_{k \in \mathcal{T}_v} \rho_t^k \qquad \forall t \in [T], v \in [V] \setminus \mathcal{V}_t
$$
\n(31)

Note  $-\beta_v \le \sum_{k \in \mathcal{T}_v} \rho_t^k$  are redundant. Combining bound (26), (28), (29), (30), and (31) and optimizing over *x*, we get an upper bound of the objective function  $\sup_{\mathbb{P}\in\mathcal{F}}\mathbb{E}_{\mathbb{P}}[c(\mathbf{x},\tilde{\mathbf{z}})]$  for any  $\mathbf{x}\in\mathcal{X}$  as ECP-C.  $\Box$ **Proof of Theorem 2.** We denote ECP-C for the case when capacity equals  $C_1$  and  $C_2$  ( $C_1 \le C_2$ ) by ECP-C1 and ECP-C2, respectively. Any feasible solution to ECP-C2 is also feasible to ECP-C1: See constraints involving *C* in ECP-C. Hence, compared to ECP-C1, the feasible region for ECP-C2 is no larger. Since the objective function of ECP-C does not depend on *C*, the optimal value of ECP-C2 is higher than that of ECP-C1.

Next, we prove the optimal value of ECP-C is less than that of ECP-U for any given value of *C*. We substitute  $a = 0$ ,  $b_v = \sum_{s \in \mathcal{T}_v} x_{v,s} e_s$ ,  $\boldsymbol{\nu} = \mathbf{0}$ ,  $\boldsymbol{y} = \boldsymbol{x}$ ,  $\boldsymbol{\rho} = \mathbf{0}$ ,  $\alpha = 0$ , and  $\boldsymbol{\beta} = \mathbf{0}$  into ECP-C: i) all the constraints are degenerated to the constraints in ECP-U; ii) the objective function becomes

$$
a + \mathbf{b}'\mathbf{\lambda} + d\left(\kappa + \gamma + \alpha + \beta'\mathbf{\lambda}\right) = \sum_{s \in [T]} e_s \sum_{v \in \mathcal{V}_s} x_{v,t} \lambda_v + d\left(\kappa + \gamma\right) = \sum_{s \in [T]} e_s f_s(\mathbf{x}, \mathbf{\lambda}) + d\left(\kappa + \gamma\right),
$$

which is the same as that of ECP-U. This means that there always exists a feasible solution in ECP-C that attains the optimal value of ECP-U. Therefore, the optimal value of ECP-C for any *C* is upper bounded by that of ECP-U.

Since the optimal value of ECP-C is monotonic in *C* and bounded by the optimal value of ECP-U, it converges to a value upper bounded by the optimal value of ECP-U as  $C \to \infty$ . **Derivation of DRO-C.** Note that  $\mathbb{E}_{\mathbb{P}}\left[\sum_{s=1}^T e_s f_s(\boldsymbol{x}, \tilde{\boldsymbol{z}})\right] = \sum_{s=1}^T e_s f_s(\boldsymbol{x}, \boldsymbol{\mu})$  for any  $\mathbb{P} \in \mathcal{F}_{\Sigma}$ . The dual of the maximization problem  $\sup_{\mathbb{P}\in\mathcal{F}_{\Sigma}} \mathbb{E}_{\mathbb{P}}\left[\max_{t\in[T]} f_t(\boldsymbol{x}, \tilde{\boldsymbol{z}})\right]$  is

$$
\inf_{\alpha,\beta,\Gamma\geq 0} \alpha + \beta'\mu + \Gamma \bullet (\mu\mu' + \Sigma) \n\text{s.t.} \quad \alpha + \beta'z + \Gamma \bullet zz' \geq f_t(x,z) \,\,\forall z \in \mathcal{Z}, \forall t \in [T].
$$
\n(32)

As  $\Gamma \succeq 0$ , the constraint in (32) can be written as  $\alpha + \beta' z + \Gamma \bullet zz' \ge \sum_{v \in \mathcal{V}_t} x_{v,t} z_v \ \forall z \in \mathcal{Z}$ , which is equivalent to  $\inf_{\mathbf{z}\in\mathcal{Z},\mathbf{U}\succeq\mathbf{z}\mathbf{z}'}\alpha+\beta'z+\Gamma\bullet\mathbf{U}-\sum_{v\in\mathcal{V}_t}x_{v,t}z_v\geq 0$ . The left-hand side of this constraint is equivalent to

$$
\inf_{z \geq 0,U} \alpha + \beta' z + \Gamma \bullet U - \sum_{v \in \mathcal{V}_t} x_{v,t} z_v
$$
\n
$$
\text{s.t.} \quad \begin{bmatrix} 1 & z' \\ z & U \end{bmatrix} \succeq 0
$$
\n
$$
\sum_{v \in \mathcal{V}_s} z_v \leq C \qquad \forall s \in [T]
$$

because  $U \succeq zz'$  is equivalent to  $\sqrt{2}$ 4  $1 z'$ *z U* 3  $\geq 0$  based on Schur's complement. Note that Slater's condition holds for the above conic linear program. The optimal value of this above conic linear program equals its dual optimal value, which is given as

$$
\sup_{\kappa,\zeta^t \ge 0, \xi^t \ge 0, y^t} \alpha - \kappa - C \sum_{s \in [T]} \zeta_s^t
$$
\n
$$
\text{s.t.} \begin{pmatrix} \kappa & \mathbf{y}^{t'} \\ \mathbf{y}^t & \mathbf{\Gamma} \end{pmatrix} \succeq \mathbf{0}
$$
\n
$$
2y_v^t = \beta_v - x_{v,t} - \xi_v^t + \sum_{s \in \mathcal{T}_v} \zeta_s^t \ \forall t \in [T], \forall v \in \mathcal{V}_t
$$
\n
$$
2y_v^t = \beta_v - \xi_v^t + \sum_{s \in \mathcal{T}_v} \zeta_s^t \qquad \forall t \in [T], \forall v \in [V] \setminus \mathcal{V}_t.
$$

As we have  $\alpha - \kappa - C \sum_{s \in [T]} \zeta_s^t \geq 0$ , (32) is equivalent to the following:

$$
\min_{\alpha,\beta,\Gamma,\atop \alpha,\beta,\Gamma,\atop \text{s.t.}} \alpha + \beta' \mu + \Gamma \bullet (\mu \mu' + \Sigma)
$$
\n
$$
\xi^{t} \geq 0, \xi^{t} \geq 0, y^{t}, \forall t \in [T]
$$
\n
$$
\text{s.t.} \qquad \begin{pmatrix} \alpha - C \sum_{s \in [T]} \zeta_{s}^{t} \ y^{t'} \\ y^{t} \end{pmatrix} \succeq 0 \qquad \forall t \in [T]
$$
\n
$$
2y_{v}^{t} = \beta_{v} - x_{v,t} - \xi_{v}^{t} + \sum_{s \in \mathcal{T}_{v}} \zeta_{s}^{t} \ \forall t \in [T], \forall v \in \mathcal{V}_{t}
$$
\n
$$
2y_{v}^{t} = \beta_{v} - \xi_{v}^{t} + \sum_{s \in \mathcal{T}_{v}} \zeta_{s}^{t} \qquad \forall t \in [T], \forall v \in [V] \setminus \mathcal{V}_{t}
$$

This formulation then gives the SDP formulation of DRO-C in §5.2.

**Proof of Proposition 4.** Note  $\mathbb{E}_{\mathbb{P}^C}\left[\sum_{v\in[V]}pu_v\tilde{z}_v\right] = \sum_{v\in[V]}pu_v\lambda_v$ . Combining it with epigraph (8) in Proposition 1 and the model in (17), we conclude the proof.  $\Box$ 

Proof of Proposition 5. Since the problems (18) and (19) are maximizing concave functions over convex sets, the objective values should be larger after optimization since they are globally optimal. Hence, the objective value is non-decreasing in Algorithm 1. Note the objective function in JPS-U is continuous with respect to pricing decision *p* and *p* is in a compact interval, and thus the objective value must be bounded. Therefore, the sequence of optimal values generated by Algorithm 1 must converge since it is monotone and bounded.  $\Box$ 

**Proof of Lemma 2.** As mentioned in §3, for any type  $v \in [V]$ , we have  $u_v \le K(\tau_v - s_v + 1)$ . For type  $v \in [V]$  with  $s_v = \tau_v$  and thus  $u_v \le K$ , (23) gives  $x_{v, \tau_v} = \min\left\{K, u_v, \max\left\{\frac{\eta X}{\sum_{v \in \mathcal{V}_t} z_v}, u_v\right\}\right\}/\eta = u_v/\eta$ , which trivially satisfies both constraints in *X*. For type  $v \in [V]$  with  $s_v < \tau_v$ , by the construction of  $x_{v,t}$ , we note that  $x_{v,t} \leq K/\eta$  (the maximum-speed constraint in *X*) and  $\sum_{\ell=s_v}^t \eta x_{v,\ell} \leq u_v$  for all  $t \in \mathcal{T}_v$ . We also note that

$$
x_{v,t} = \min\left\{K, u_v - \sum_{\ell=s_v}^{t-1} \eta x_{v,\ell}, \max\left\{\frac{\eta X}{\sum_{v \in \mathcal{V}_t} z_v}, u_v - \sum_{\ell=s_v}^{t-1} \eta x_{v,\ell} - K(\tau_v - t)\right\}\right\}/\eta
$$
  

$$
\geq \min\left\{K, u_v - \sum_{\ell=s_v}^{t-1} \eta x_{v,\ell}, \frac{\eta X}{\sum_{v \in \mathcal{V}_t} z_v}\right\}/\eta \geq 0,
$$

where the last inequality follows from  $\sum_{\ell=s_v}^{t-1} \eta x_{v,\ell} \leq u_v$  and  $X \geq 0$ . Thus, the second constraint in *X* holds. We next show the first constraint in *X* holds by proving  $\sum_{\ell=s_v}^t \eta x_{v,\ell} \ge u_v - K(\tau_v - t)$  for all  $t \in \mathcal{T}_v$  via induction.

In the first period  $t = s_v$ , (23) gives  $x_{v,s_v} = \min\left\{K, u_v, \max\left\{\frac{\eta X}{\sum_{v \in \mathcal{V}_{s_v}} z_v}, u_v - K(\tau_v - s_v)\right\}\right\}/\eta \ge$  $\min\{K, u_v, u_v - K(\tau_v - s_v)\}/\eta = (u_v - K(\tau_v - s_v))/\eta$ : The last equality is due to  $\min\{K, u_v\} \ge u_v - K(\tau_v - s_v)/\eta$  $s_v$ ), which in return results from  $K(\tau_v - s_v) \geq 0$  and  $K(\tau_v - s_v + 1) \geq u_v$ . Thus, the hypothesis holds for  $t = s_v$ . Suppose the hypothesis holds for all  $t \leq t'-1$ , we next prove this holds for  $t = t' \in \mathcal{T}_v$ . Note that

$$
\eta x_{v,t'} \ge \min \left\{ K, u_v - \sum_{\ell=s_v}^{t'-1} \eta x_{v,\ell}, u_v - \sum_{\ell=s_v}^{t'-1} \eta x_{v,\ell} - K(\tau_v - t') \right\} = u_v - \sum_{\ell=s_v}^{t'-1} \eta x_{v,\ell} - K(\tau_v - t'),
$$

where the equality is because  $\sum_{\ell=s_v}^{t'-1} \eta x_{v,\ell} \ge u_v - K(\tau_v - t' + 1)$  by the induction hypothesis for period  $t'-1$ . This completes the induction of  $\sum_{\ell=s_v}^{t} \eta x_{v,\ell} \ge u_v - K(\tau_v - t)$ .

When  $t = \tau_v$ , we obtain  $\sum_{\ell=s_v}^{\tau_v} \eta x_{v,\ell} = u_v$  by noting that  $\sum_{\ell=s_v}^{\tau_v} \eta x_{v,\ell} \leq u_v$  and  $\sum_{\ell=s_v}^{\tau_v} \eta x_{v,\ell} \geq u_v$  from the above. In sum, the EQC policy defined in  $(23)$  is feasible.  $\Box$ 

#### Appendix B: Data Calibration

Demand charge calibration.  $\hat{d}_t$  is obtained by dividing the sum of three demand charges in a month by 30 (i.e., scaling to those of a day): all-period demand charge (termed "facility-related demand charge" in the tariff structure, and applicable to all periods throughout the horizon at  $13.94/\text{kW}$ ; peak-hour demand charge (applicable only to on-peak hours at \$16.20/kW); and mid-peak demand charge (applicable only to mid-peak hours at \$4.95/kW).

Preprocess EV charging data. We preprocess the data as follows: We remove those charging events where either the plug-in duration (the time interval between start time and end time) is more than 4 hours (as most are less than 4 hours), or the charging power (total electricity charged divided by plug-in duration) is larger than 43 kW, or the total electricity charged is either less than 0.5 kWh or greater than 62.5 kWh. In the end, we retain 93.53% of the original charging events and obtain 101,707 charging events in total, i.e., 278.65 daily charging events on average. Note that the sample size for calibrating each type is 365 as each day has a data point.

Long-duration charging problem setup. The problem setup is the same as the short-duration charging setup, except we use the data for fast chargers, where the maximal charging rate is 22 kW. In this case, the stay duration is longer, and we assume the longest stay duration is 8 hours, i.e., *L* = 32, greater than that of 70% of the charging events in the data. This duration is also reasonable for workplace charging. We set the maximal charging requirement to be 30 kWh, which is larger than that of 98% of the charging events in the data. In this case, the expected number of EVs each day is 196.91, and the number of types *V* is 85*,* 536, among which 21,426 have positive arrival rates. (Note that if the longest stay duration is a whole day, i.e.,  $L = 96$ , we have 269,856 types, among which 33,177 have positive arrival rates. It is much more computationally intensive.) We plot the average number of EVs at the station in Figure 8.

#### Appendix C: Details of computation results on small, six-period instances

**Experiment settings.** The longest stay duration for all EVs is 3 periods. We set  $U = 6$  kWh and  $u \in [U]$ ;  $K = 3$  kW/period; and  $\eta = 0.9$ . We set the unit energy charge in period *t* to be  $\hat{e}_{16t}$  (see §7.1). Let us denote the calibrated arrival rates for the 96-period instances by  $\hat{\lambda}$ . We then set the arrival rate for the 6-period



Figure 8 Average number of EVs at the charging station for the long-duration charging case

instances to be  $\hat{\lambda}$  taken every 16 periods, and multiply them by 16 times in order to have a comparable number of charging events in the horizon as in the 96-period case, i.e., the arrival rate for type  $(s, \tau, u)$  equals  $16\lambda_{16s,16s+(\tau-s),3u}$ . We also vary the number of chargers, *C*, from the set  $\{12, 10, 8\}$ , for which the service level is 98.69%, 95.65%, and 88.55%, respectively.

Computation and evaluation. For DRO-C, we generate 5*,* 000 samples and use the sample mean and covariance as  $\mu$  and  $\Sigma$  in the ambiguity set (15), respectively. We use SAA with a 5,000 sample size, which already converges, and implement SAA five times and take the average performance of these five solutions. We use the interior point method of SAA (converging within 0.5 seconds for all values of *C* and *d* considered, see Table 12), faster than the L-shaped method with multi-cut, which is faster than this method with single cut. Figure 9(a) and (b) plot for the L-shaped method with single cut both the lower and upper bounds of the objective value and CPU time in each iteration; those of multi-cut version are plotted in Figure 9(c) and (d). As seen, the multi-cut version converges within four seconds, while the single cut version cannot converge within 12 hours and converges very slowly after several iterations.

Performance of ECP-C, SAA, and DRO-C. Table 12 lists for all three approaches under different values of *C* and *d* the performance evaluated on 10,000 sample paths, including the mean and standard deviation of the total cost, expected ratio of the demand charge and the total cost, and CPU time. Among all three approaches, SAA has the lowest mean cost, which is statistically significant given the small standard error of the mean cost under any approach. This is expected as it converges to the optimal solution.

DRO-C has the second-lowest mean total cost, lower than ECP-C. This may be because DRO-C gives a sharper upper bound of the SP model than ECP-C does on *small instances*. However, the performance of DRO-C should deteriorate as the problem size increases. We illustrate this by comparing the upper bound of  $\mathbb{E}_{\mathbb{P}}\left[\max_{v\in[V]}\tilde{z}_v\right]$  computed using both DRO and ECP where  $\tilde{z}_v \sim Poisson(\lambda_v)$  and *V* is the dimension of the random vector, which represents the problem size. We set  $\lambda_v$  equal to 7 and 75 so that

$\,C$	$\boldsymbol{d}$	Method	Total cost		$\mathbb{E}\left\{\text{Demand charge}\right\}$	CPU
			Mean	$\operatorname{Std}$	$/total cost$ }	Time(s)
		$ECP-C$	6.00	1.76	0.43	0.08
	0.23	<b>SAA</b>	5.96	1.74	0.45	0.50
12		DRO-C	5.97	1.75	0.44	1354.95
		$ECP-C$	8.64	2.53	0.60	0.06
	0.46	<b>SAA</b>	8.57	2.49	0.61	0.49
		$DRO-C$	8.60	2.51	0.60	1223.42
		$ECP-C$	13.90	4.08	0.75	0.06
	0.93	SAA	13.74	4.01	0.75	0.50
		DRO-C	13.80	$4.05\,$	0.75	1039.48
		$ECP-C$	5.75	$1.52\,$	0.44	0.05
	0.23	<b>SAA</b>	5.72	1.52	0.45	0.49
		$DRO-C$	5.73	$1.52\,$	0.44	1365.50
	0.46	ECP-C	8.28	2.19	0.60	0.06
10		SAA	8.23	2.18	0.61	0.48
		DRO-C	8.25	2.19	0.60	1304.83
		$ECP-C$	13.49	$3.56\,$	0.75	0.05
	0.93	<b>SAA</b>	13.20	$3.51\,$	0.75	0.50
		DRO-C	13.25	3.53	0.75	1027.80
		$ECP-C$	5.41	1.25	0.43	0.06
	0.23	<b>SAA</b>	$5.22\,$	1.22	0.45	0.43
		$DRO-C$	5.23	1.22	0.44	1217.52
		$ECP-C$	8.08	1.84	0.61	0.05
8	0.46	<b>SAA</b>	7.51	1.75	0.60	0.43
		$DRO-C$	7.54	1.76	0.60	1202.28
		$ECP-C$	12.99	$2.95\,$	0.76	0.05
	0.93	<b>SAA</b>	12.04	2.83	0.75	0.42
		DRO-C	12.11	2.83	0.75	1114.59

Table 12 Performance of ECP-C, SAA, and DRO-C on small, six-period instances

 $\mathbb{E}_{\mathbb{P}}\left[\max_{v\in[V]}\tilde{z}_v\right]$  is close to the expected peak load under the optimal solution in the 6-period and 96-period instance (i.e.,  $\mathbb{E}_{\mathbb{P}}\left[\max_{t\in[6]} f_t(\boldsymbol{x}^*, \tilde{\boldsymbol{z}})\right]$  and  $\mathbb{E}_{\mathbb{P}}\left[\max_{t\in[96]} f_t(\boldsymbol{x}^*, \tilde{\boldsymbol{z}})\right]$ ), respectively. Figure 10 plots the upper bound of  $\mathbb{E}_{\mathbb{P}}\left[\max_{v\in[V]}\tilde{z}_v\right]$  obtained by DRO and ECP compared with the true value of  $\mathbb{E}_{\mathbb{P}}\left[\max_{v\in[V]}\tilde{z}_v\right]$ obtained via Monte-Carlo estimation. As seen, ECP bound becomes much sharper than DRO for large *V* (see Figure 10b). Therefore, even if DRO-C can be solved for a large instance, its solution may not be as good as ECP-C.

In addition, ECP-C is around eight times faster than SAA, both of which are thousands of times faster than DRO-C because DRO-C solves an SDP.

#### Appendix D: Performance of SAA

Figure 11 plots the performance of SAA at different sample sizes, where Figure  $11(a)$  shows that at the largest sample size considered (i.e., 8,000), SAA has the lowest mean and standard deviation of the total cost, though also the longest computation time (in Figure 11(b)).

#### Figure 9 Comparison of single cut and multi-cut L-shaped methods on six-period instances  $(C = 10, d = 0.46)$



Figure  $10$   $\:$  Upper bound of  $\mathbb{E}_{\mathbb{P}}\left[\max_{v\in[V]}\tilde{z}_v\right]$  obtained by DRO and ECP where  $\tilde{z}_v\sim Poisson(\lambda_v)$ 



*Notes.* In (b), note that DRO for the cases when the dimension  $V > 46$  cannot be solved to the desired accuracy and thus is not shown.

### Appendix E: Rolling horizon implementation of LP and ECP

For notational convenience, for any  $t \in [T]$ , we write  $[V] = [V]^{t+} \cup [V]^{t-}$ , where  $[V]^{t+} = \{v \in [V] : s_v \ge t\}$  and  $[V]^{t-} = \{v \in [V] : s_v \le t-1\}$ , denoting the customer types that will arrive from period *t* onward and those that have arrived by period  $t-1$ , respectively. Recall from §3 that  $V_{t}$  is the set of customer types at the

#### Figure 11 Performance of SAA on 96-period instances with different sample sizes at  $C = 30$



(a) Mean and standard deviation of total cost (b) Total CPU time in seconds *Notes.* In (a), the number in the parenthesis of the legend represents the sample size.

station in period t'. For any  $t' \ge t \in [T]$ , we denote  $\mathcal{V}_{t'}^{t+} = \{v \in \mathcal{V}_{t'} : s_v \ge t\}$  and  $\mathcal{V}_{t'}^{t-} = \{v \in \mathcal{V}_{t'} : s_v \le t-1\}$  as those from the set  $V_{t}$  that arrive at the station between period *t* and period *t'* and that arrive by period  $t-1$ , respectively. For each  $t \in [T]$ , we denote the implemented decision  $\hat{x}_{v,s}$  for  $s \leq t-1$  and  $v \in [V]^{t-1}$ given uncertainty realization  $\hat{z}_v$  for all  $v \in [V]^{t-}$ . Given  $\hat{x}_{v,s}$ , the feasible set for the decision variable  $x_t =$  $(x_{v,s})_{s>t,v\in\mathcal{V}_s}$  is given as follows:

$$
\mathcal{X}_t \triangleq \left\{ \boldsymbol{x}_t : \sum_{s=t}^{\tau_v} x_{v,s} \le K/\eta \sum_{s=s_v}^{t-1} \hat{x}_{v,s} \ \forall v \in [V] \text{ with } \tau_v \le t \right\}.
$$
\n(33)

That is,  $x_t \in \mathcal{X}_t$ . Note we associate variables defined in the main text with subscript t to indicate these variables update in each period  $t$  and omit their definition when no confusion arises. Let  $\hat{f}_t^{\text{max}}$  denote the  $\max_{\mathbf{x} \in [t]} \{f_s(\hat{\mathbf{x}}, \hat{\mathbf{z}})\}\text{, where we set }\hat{f}_0^{\max} = 0.$ 

#### E.1. RH-LP

The RH-LP policy can be regarded as a "predict-then-optimize" approach, which is implemented in Algorithm 3. In each period *t*, we first solve a deterministic LP problem over the remaining time horizon, i.e., from period t to T. In particular, we replace in the corresponding SP model the past arrivals  $(\tilde{z}_v)_{v \in |V|^{t-}}$  by realizations  $\hat{z}_v$  and future arrivals  $(\tilde{z}_v)_{v \in [V]^{\tau+}}$  by their mean values  $\mu_v = \mathbb{E}_{\mathbb{P}}[\tilde{z}_v]$ , in (34). Note that in (34) the energy charge from period t to T has two terms:  $\sum_{v \in V_s^t} x_{v,s} \hat{z}_v$ , representing energy charge of vehicles that have arrived before *t* but have not finished charging yet, and  $\sum_{v \in \mathcal{V}_{s}^{t}} x_{v,s} \mu_v$ , representing energy charge of vehicles that have not arrived yet. The two constraints in (34) are the linearization of the max operator in the peak load max  $\left\{\hat{f}_{t-1}^{\max}, \max_{s \in \{t, ..., T\}} \left\{ \sum_{v \in \mathcal{V}_s^{t-}} x_{v,s} \hat{z}_v + \sum_{v \in \mathcal{V}_s^{t+}} x_{v,s} \mu_v \right\} \right\}$ 

We then implement the decisions only in the current period (step 2) and finally update the implemented decisions  $\hat{x}_{v,t}$ , the feasible set  $\mathcal{X}_{t+1}$  using (33), and  $\hat{f}_t^{\max}$  (step 3).

#### E.2. RH-ECP

We develop Algorithm 4 to implement RH-ECP. Algorithm 4 is exactly the same as Algorithm 3 except in the first step Algorithm 4 resolves in each period *t* an ECP problem over the remaining time horizon (from

## Algorithm 3: Rolling Horizon LP (RH-LP)

for  $t = 1, 2, 3, ..., T$  do 1. Given uncertainty realization  $\hat{z}_v$  for all  $v \in [V]^{t-}$ , solve the following (deterministic) LP:  $\min_{\mathbf{x}_t \in \mathcal{X}_t} d\gamma + \sum_{t=1}^T$ *s*=*t es*  $\sqrt{ }$  $\left( \right)$  $v \in \mathcal{V}_s^{t-1}$  $x_{v,s}\hat{z}_v + \sum$  $v \in \mathcal{V}_s^{t+1}$  $x_{v,s}\mu_v$ 1 A s.t.  $\hat{f}_{t-1}^{\max} \leq \gamma$  $\sum$  $v \in \mathcal{V}_s^t$ <sup>-</sup>  $x_{v,s}\hat{z}_v + \sum$  $v \in \mathcal{V}_s^{t+1}$ *xv,sµ<sup>v</sup>* 8*s t* (34) 2. Observe uncertainty realizations  $\hat{z}_v$  with  $s_v = t$  and implement  $x_{v,t}$  with  $v \in \mathcal{V}_t$ . 3. Update  $\hat{x}_{v,t} \leftarrow x_{v,t}$  for  $v \in \mathcal{V}_t$ ,  $\mathcal{X}_{t+1}$  using (33), and  $\hat{f}_t^{\max} \leftarrow \max\left\{\hat{f}_{t-1}^{\max}, \sum_{v \in \mathcal{V}_t} \hat{x}_{v,t} \hat{z}_v\right\}$ end

period *t* to *T*) similar to the static version of the ECP. To obtain this ECP formulation in each period *t*, note that given the realized arrivals  $\tilde{z}_v = \hat{z}_v$  for all  $v \in [V]^{t-}$ , we update the support  $\mathcal{Z}_t$  for the future arrivals  $\tilde{\mathbf{z}}_t \triangleq (\tilde{z}_v)_{v \in [V]^{\text{t+}}}$ , which differs from (1) and is given as follows:

$$
\mathcal{Z}_t \triangleq \left\{ \mathbf{z}_t \geq \mathbf{0} : \sum_{v \in \mathcal{V}_s^{t+}} z_v \leq C - \sum_{v \in \mathcal{V}_s^{t-}} \hat{z}_v, \forall s \in \{t, ..., T\} \right\}.
$$

We then update the ambiguity set  $\mathcal{F}_t$  by updating  $\mathcal{Z}_t$  and  $[V]^{t+}$  in the definition of  $\mathcal{F}_t$  as follows:

$$
\mathcal{F}_t \triangleq \left\{ \mathbb{P} \in \mathcal{P}_0 \left( \mathbb{R}^{|[V]^{t+}|} \right) \middle| \begin{array}{c} \tilde{z}_t \triangleq (\tilde{z}_v)_{v \in [V]^{t+}} \sim \mathbb{P}_t \\ \ln \mathbb{E}_{\mathbb{P}_t} \left[ \exp \left( \sum_{v \in [V]^{t+}} \theta_v \tilde{z}_v \right) \right] \leq \sum_{v \in [V]^{t+}} \lambda_v (e^{\theta_v} - 1), \quad \forall \theta \geq 0 \\ \mathbb{P}_t \left[ \tilde{z}_t \in \mathcal{Z}_t \right] = 1 \end{array} \right\}.
$$

The entropic dominance constraints (the second constraint in  $\mathcal{F}_t$ ) are the same as those in  $\mathcal F$  except restricting to future arrivals. Hence, we can still interpret these constraints as the truncation on independent Poisson variables in the corresponding uncapacitated case. Similar to the static ECP, we can write  $\mathcal{F}_t = \mathcal{F}_t^1 \cap \mathcal{F}_t^2$ where

$$
\mathcal{F}_t^1 \triangleq \left\{ \mathbb{P}_t \in \mathcal{P}_0(\mathbb{R}^{|[V]^{t+}|}) \middle| \ \tilde{z}_t \triangleq (\tilde{z}_v)_{v \in [V]^{t+}} \sim \mathbb{P}_t; \ln \mathbb{E}_{\mathbb{P}_t} \left[ \exp \left( \sum_{v \in [V]^{t+}} \theta_v \tilde{z}_v \right) \right] \leq \sum_{v \in [V]^{t+}} \lambda_v (e^{\theta_v} - 1), \ \ \forall \theta \geq 0 \right\}
$$
\n
$$
\mathcal{F}_t^2 \triangleq \left\{ \mathbb{P}_t \in \mathcal{P}_0(\mathbb{R}^{|[V]^{t+}|}) \middle| \ \tilde{z}_t \triangleq (\tilde{z}_v)_{v \in [V]^{t+}} \sim \mathbb{P}_t; \mathbb{E}_{\mathbb{P}_t} [\tilde{z}_t] \leq \lambda_t; \mathbb{P}_t [\tilde{z}_t \in \mathcal{Z}_t] = 1 \right\}
$$

Similar to DRO-Ent, we aim to solve the following DRO problem in each period:

$$
\min_{\mathbf{x}_t \in \mathcal{X}_t} \sup_{\mathbb{P}_t \in \mathcal{F}_t} \mathbb{E}_{\mathbb{P}_t} \left[ d \max \left\{ \hat{f}_{t-1}^{\max}, \max_{s \in \{t, \dots, T\}} \left\{ \sum_{v \in \mathcal{V}_s^t} x_{v,s} \hat{z}_v + \sum_{v \in \mathcal{V}_s^{t+}} x_{v,s} \tilde{z}_v \right\} \right\} + \sum_{s=t}^T e_s \left( \sum_{v \in \mathcal{V}_s^t} x_{v,s} \hat{z}_v + \sum_{v \in \mathcal{V}_s^{t+}} x_{v,s} \tilde{z}_v \right) \right],
$$
\n
$$
(35)
$$

where (35) differs from DRO-Ent in the following: (i) the objective function is from period  $t$  to  $T$ ; (ii) the feasible set  $\mathcal{X}_t$  depends on the implemented decision  $\hat{x}_{v,s}$  for  $s \leq t-1$  and  $v \in [V]^{t-}$ ; (iii) past arrivals

#### Algorithm 4: Rolling Horizon ECP (RH-ECP)

for  $t = 1, 2, 3, ..., T$  do 1. Given uncertainty realization  $\hat{z}_v$  for all  $v \in [V]^{t-}$ , solve min  $x_t \in X_t, a, b \ge 0, \nu \ge 0, y^0 \ge 0, y \ge 0, \mu > 0, \kappa, \gamma, \alpha, \beta \ge 0, \xi, \zeta, \rho \ge 0$ *d*  $\sqrt{ }$  $\left( \kappa + \gamma + \alpha + \sum_{\alpha} \right)$  $v \in [V]$ <sup>t+</sup>  $\beta_v \lambda_v$ 1  $+ a$  $+$   $\sum$  $v \in [V]^{t+1}$  $b_v\lambda_v+\sum_{i=1}^T$ *s*=*t*  $e_s$ <sup>y</sup>  $v \in \mathcal{V}_s^{t-1}$  $x_{v,s}\hat{z}_v$ s.t.  $y^0 \le \gamma$  $\overline{\phantom{0}}$  $v \in \mathcal{V}_s^{t-1}$  $y_{v,s}\hat{z}_v + \sum$  $v \in \mathcal{V}_s^{t+1}$ *yv,s<sup>v</sup>* 8*s t*  $(\xi_{v,s}, \mu, y_{v,s}) \in \mathcal{K}_{\exp}$  8*y*<sup>t</sup> + 8*s* 2*t*, *v*  $\in \mathcal{V}_{s}^{t+1}$  $( \zeta v, s, \mu, y_{v,s}) \subset \mathcal{K}_{\exp}$   $\qquad \qquad \searrow \qquad \searrow \qquad \searrow \qquad \searrow \qquad \searrow \qquad \searrow \qquad \searrow$  $\left(\zeta_s, \mu, -\kappa + \sum\right)$  $v \in \mathcal{V}_s^{t+1}$  $\lambda_v(\xi_{v,s} - y_{v,s} - \mu)$ 1  $\begin{cases} \in \mathcal{K}_{\rm exp} \quad & \forall s \geq t \end{cases}$ X *T s*=*t*  $\zeta_s \leq \mu$  $\hat{f}_{t-1}^{\max} - y^0 \leq \alpha$ X *T*  $k = t$  $\sqrt{ }$  $\Big(C-\sum\limits_{i=1}^{n}$  $v \in \mathcal{V}_k^{t-1}$  $\hat{z}_v$ 1  $\left| \rho_s^k \leq \alpha - \sum \right|$  $v \in \mathcal{V}_s^{t-1}$  $(x_{v,s} - y_{v,s})\hat{z}_v \ \forall s \ge t$  $x_{v,s} - y_{v,s} - \beta_v \leq \sum$  $k$  $\in$   $\mathcal{T}_v$  $\rho^k_s$ *s*  $\forall s \ge t, \forall v \in [V]^{t+1}$ X *T s*=*t*  $\sqrt{ }$  $\Big(C-\sum\limits_{i=1}^{n}$  $v \in \mathcal{V}_s^t$ <sup>-</sup>  $\hat{z}_v$  $\sqrt{ }$  $\nu_s \leq a$  $\sqrt{ }$  $s \in \mathcal{T}_v$  $x_{v,s}e_s - b_v \leq \sum$  $s \in \mathcal{T}_v$  $\nu_s$   $\forall v \in [V]^{t+1}$ (36)

2. Observe uncertainty realizations  $\hat{z}_v$  with  $s_v = t$  and implement  $x_{v,t}$  with  $v \in \mathcal{V}_t$ .

3. Update  $\hat{x}_{v,t} \leftarrow x_{v,t}$  for  $v \in \mathcal{V}_t$ ,  $\mathcal{X}_{t+1}$  using (33), and  $\hat{f}_t^{\max} \leftarrow \max\left\{\hat{f}_{t-1}^{\max}, \sum_{v \in \mathcal{V}_t} \hat{x}_{v,t} \hat{z}_v\right\}$ . end

 $(\tilde{z}_v)_{v \in [V]^t}$  are replaced by their realizations  $(\hat{z}_v)_{v \in [V]^t}$ ; and (iv) the peak load is the maximum of historical peak load  $\hat{f}_{t-1}^{\text{max}}$  and future peak load from period t to T.

We then develop a tractable upper bound of (35) in (36), in Algorithm 4, which then becomes the upper bound of the corresponding SP problem. We omit the derivation as it is similar to that in the proof of Proposition 3.